Impact of home food production on nutritional blindness, stunting, wasting, underweight and mortality in children: a systematic review and meta-analysis of controlled trials

Chizoba Bassey^{a*}, Harriet Crooks^b, Katherine Paterson^c, Rachel Ball^c, Kristoffer Howell^c, Iona Humphries-Cuff^c, Kirsty Gaffigan^c, Nitya Rao^d, Jennifer A. Whitty^{a, e}, Lee Hooper^a

^aNorwich Medical School, University of East Anglia, Norwich, United Kingdom; ^bStudent Services, University of East Anglia, Norwich, United Kingdom; ^cDepartment of Nutrition and Dietetics, Norfolk and Norwich University Hospitals NHS Foundation Trust, Norwich, United Kingdom; ^dSchool of Developmental Studies, University of East Anglia, Norwich, United Kingdom; ^e National Institute for Health Research (NIHR) Applied Research Collaboration (ARC) East of England, United Kingdom

*Corresponding author: Chizoba Bassey, <u>c.nwabichie@uea.ac.uk</u>, Norwich Medical School, University of East Anglia, Norwich NR4 7TJ, United Kingdom.

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Vitamin A deficiency is highly prevalent and remains the major cause of nutritional blindness in children in low-and middle-income countries, despite supplementation programmes. Xeropthalmia (severe drying and thickening of the conjunctiva) is caused by vitamin A deficiency and leads to irreversible blindness. Vitamin A supplementation programmes effectively reduce vitamin A deficiency but many rural children are not reached. Home food production may help prevent rural children's vitamin A deficiency. We aimed to systematically review trials assessing effects of home food production (also called homestead food production and agricultural interventions) on xeropthalmia, nightblindness, stunting, wasting, underweight and mortality (primary outcomes).. We searched Medline, Embase, Scopus, Cochrane CENTRAL and trials registers to February 2019. Inclusion of studies, data extraction and risk of bias were assessed independently in duplicate. Random-effects metaanalysis, sensitivity analyses, subgrouping and GRADE were used. We included 16 trials randomizing 2498 children, none reported xerophthalmia, night-blindness or mortality. Home food production may slightly reduce stunting (mean difference (MD) 0.13 (z-score), 95% CI 0.01 to 0.24), wasting (MD 0.05 (z-score), 95% CI -0.04 to 0.14) and underweight (MD 0.07 (z-score), 95% CI -0.01 to 0.15) in young children (all GRADE low-consistency evidence), and increase dietary diversity (standardized mean difference (SMD) 0.24, 95% CI 0.15 to 0.34). Home food production may usefully complement vitamin A supplementation for rural children. Large, long-duration trials with good randomization, allocation concealment and correct adjustment for clustering are needed to assess effectiveness of home food production on nutritional blindness in young children.

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Introduction

Approximately 250 million preschool children are vitamin A deficient (WHO, 2018). Vitamin A deficiency is the main global cause of preventable childhood blindness with about 2.8 million preschool-age children at risk of blindness (WHO, 2009). Vitamin A deficiency also increases the risk of mortality from other childhood diseases such as diarrhea and measles and plays a significant role in normal immune function (UNICEF, 2018). Vitamin A deficiency remains one of the most prevalent micronutrient deficiencies globally, being most predominant in low and middle income countries (Bahreynian et al., 2017).

Nutritional blindness in children is caused by vitamin A deficiency and manifests as xerophthalmia which is an array of ocular signs and symptoms. Xerophthalmia presents as night blindness, Bitot's spots, conjunctival xerosis, corneal xerosis, corneal ulcer, corneal scarring and keratomalacia. It mainly affects children under 5 years of age, with the greatest concentration in preschool children. Once the cornea becomes ulcerated and melts away, blindness becomes irreversible (Gilbert, 2013). Xerophthalmia is predominantly caused by an insufficient intake of vitamin A (Akhtar et al., 2013).

One of the major causes of vitamin A deficiency is insufficient dietary intake of foods rich in vitamin A (including retinol, retinal, retinoic acid, and pro-vitamin A carotenoids such as beta-carotene) (Wirth et al., 2017). Vitamin A deficiency is more prevalent in children living in rural areas (Sherwin et al., 2012; Schemann et al., 2007; Dole et al., 2009; Hanson et al., 2016). A cross sectional study carried out in the Democratic Republic of Congo by Samba et al. (2006) revealed that larger quantities of vitamin A-rich foods were consumed in urban than rural areas. Bitot's spots were significantly more prevalent in children of lower socioeconomic status (P < 0.001) in a cross-sectional survey from rural India by Arlappa et al. (2011) and there was a gross deficiency of vitamin A-rich food in rural areas.

Programmes of high dose vitamin A supplementation occur in 82 countries, and involve supplementing children aged 6 to 59 months with retinol twice yearly (UNICEF, 2018). According to a systematic review and meta-analysis of randomized controlled trials carried out by Imdad et al in 2017, vitamin A supplementation caused a 12% reduction in all-cause mortality (Risk Ratio (RR) 0.88, 95% CI 0.83 – 0.93). It also reduced the risk of diarrhea (RR 0.85, 95%Cl 0.82 – 0.87), measles (RR 0.50, 95%Cl 0.37 – 0.67), night blindness (RR 0.32, 95% Cl 0.21 – 0.50) and Bitot's spots (RR 0.42 95% Cl 0.33 – 0.53) in children below 5 years of age. Although supplementation programmes are clearly beneficial, programme coverage can be low (Semba et al., 2008). In 2016, vitamin A supplementation programmes reached only 64% of targeted children (UNICEF, 2018). Of the 82 priority countries for vitamin A supplementation, only 57 achieved two-dose coverage (UNICEF, 2018). Coverage is not equitable. Thapa (2008) analyzed data from the 2006 Nepal Demographic and Health Survey, reporting that vitamin A supplementation bypassed the poorest of the poor, illiterate mothers and rural inhabitants. Nguyen et al. (2012) examined associations between vitamin A supplementation program coverage and socio-demographic factors in Nepal. Ironically, they found that children living in rural areas, who are more susceptible to vitamin A deficiency, were less likely to receive vitamin A supplements than children living in urban areas.

Food based approaches may have the potential for achieving sustainability in controlling vitamin A deficiency and can complement vitamin A supplementation programmes (Chakravarty, 2000). Food based approaches include point-of-use fortification with micro-nutrient powders, home food production of vitamin A-rich crops and fortification of staple foods such as oil, wheat flour, sugar with vitamin A (Chakravarty, 2000). Although point-of-use fortification appears easy and efficacious in reducing vitamin A deficiency, achieving high coverage, adherence and appropriate doses (as excessive doses can lead to toxicity) are key challenges that hamper its utility (Dhillon et al. 2017). Food fortification tends to be less

useful in resource-poor settings as poorer people make food choices based on price rather than quality (Dary and Mora 2002).

A systematic review assessing the effects of home food production on key outcomes of vitamin A deficiency, including nutritional blindness, in children would help to understand the utility of this approach. Existing systematic reviews (Girard et al., 2012; Masset et al., 2012) that assess effects on vitamin A status and nutritional outcomes in children are outdated and do not assess the impact of home food production on nutritional blindness. This review aims to systematically assess the effectiveness of home food production on nutritional blindness and anthropometric measures in children.

Materials and methods

Our systematic review methodology was based on the Cochrane Handbook (Higgins et al., 2019) and reported using PRISMA guidelines (Moher et al., 2015).

Search strategy

A protocol was developed, registered with PROSPERO (CRD42019126455) and used to ensure methodological rigor. Medline Ovid, Embase Ovid, Scopus, Cochrane Central Register of Controlled Trials and The World Health Organization International Clinical Trial Registry Platform were searched from inception to the 1st of February 2019. A complex search strategy using text words, index terms, truncation and Boolean operators was developed using the following framework: ((children or women) AND (home gardening) AND (RCT OR CCT)) (Lefebvre et al., 2019). See Appendices 1 – 5 in the Supplementary Materials for full search strategies.

Inclusion/exclusion criteria

We included controlled clinical trials (CCTs) (studies with a concurrent intervention and control arm) and randomized controlled trials (RCTs) of at least one year duration. Participants were women of childbearing age. Interventions included were provision of seedlings and/or training in the planting of vitamin A- rich foods and other crops on a piece of land attached to the home or near home primarily for household consumption, training in and/or distribution of chicks for home rearing and consumption with or without behavioral change interventions. The comparator needed to be a non-intervention group or an alternative approach to home gardening. Studies without a comparator and before/after studies were excluded, as were studies that assessed commercial or school farming. Primary outcomes were night blindness, xerophthalmia, all-cause mortality, stunting, wasting, underweight (all assessed in children). Secondary outcomes included family income, children's serum retinol levels, children's dietary diversity, and the cost of intervention.

Data extraction and synthesis

Search results were uploaded to Covidence software (Covidence, 2019). Two independent reviewers screened titles and abstracts, and then full texts using the inclusion/exclusion criteria. Where inclusion of a title/abstract was unclear, or appeared likely, the full text was obtained for full assessment. Conflicts were settled through discussion. A data extraction form bespoke to this review was developed, tested and used in data extraction. Data were extracted and risk of bias assessed by two independent reviewers and conflicts resolved by a third reviewer. The Cochrane risk of bias tool (Higgins et al., 2011) was used in assessing risk of bias in this systematic review, studies were assessed for selection bias, performance bias, detection bias, attrition bias and reporting bias (Higgins et al., 2011). A study outcome

was judged to be at low risk of bias if that outcome was at low risk of bias for all domains, otherwise it was judged at high risk of bias (Higgins et al. 2011).

Data were tabulated, meta-analysis and narrative synthesis were used. Random effects metaanalysis (in Review Manager version 5.3 (Review Manager, 2014)) was the default analysis as we assumed included studies had related intervention effects and were similar enough to pool but displayed slightly different effect estimates (Deeks et al., 2019). Statistical heterogeneity was examined using I^2 (Deeks et al., 2019), assuming that $I^2 >75\%$ represented important heterogeneity (Deeks et al., 2019). We sub-grouped by methodology – some studies (appropriately) adjusted data for clustering, some adjusted for clustering and other factors such as age, sex, education (important and appropriate where there are few clusters), and some were unadjusted (which is statistically inappropriate). We subgrouped by adjustments to prevent overestimation of effects. Our main analysis included data adjusted for clustering and other factors, to reduce effects of confounding factors resulting from randomization in large clusters.

We ran sensitivity analyses using the fixed effects model of meta-analysis (Deeks et al., 2019). The following subgroups were analyzed for primary outcomes:

- Studies that focused on planting of fruits, vegetables and rearing of chickens versus only planting of fruits and vegetables.
- Studies with a duration of $12 \le 24$ months versus > 24 months
- Studies conducted in Africa versus studies conducted in Asia.

Small study bias was assessed by comparing random- and fixed-effects meta-analysis results. The quality and consistency of the findings in this review were assessed and reported using GRADE assessment in GRADEpro GDT (Schunemann et al., 2019, GRADE pro G.D.T, 2015).

Results

In total, 7021 titles and abstracts were retrieved and uploaded to Covidence, 1623 were eliminated as duplicates, and 5398 titles and abstracts were screened using the inclusion and exclusion criteria. One hundred and fifteen full text papers were collected of which 92 were excluded. Twenty-three papers were included and merged into 16 individual included studies (Figure 1). Seven further studies were found eligible but are on-going, with no published outcomes (Supplementary Table 1).

Included studies were conducted in Africa (nine studies) and Asia (seven studies). Eleven studies had a duration of 12 to \leq 24 months and five studies had a duration of > 24 months. Studies were cluster randomized trials (ten studies), individually randomized trials (one study), and controlled clinical trials (five studies). Gelli et al., 2018 reported their data grouping children into 6 – 24 months and 36 – 72 months of age, so results were reported separately for these two groups. See Table 1 for characteristics of all included studies. Types of intervention in the included studies ranged from training in setting up home gardens, rearing of chicks and other domestic animals, cooking sessions, offer of loans to set up home gardens, training in selling of surplus produce, nutrition education, distribution of seedlings, chicks, orange sweet potato, and other planting materials. Most of the comparator groups had no intervention (Table 1).

Six of the sixteen trials were at low risk of bias from randomization, none from allocation concealment, 3 from blinding of participants and personnel, and 3 from blinding of outcome assessment. All included studies were judged at high risk of bias (See Figure 2 for risk of bias details trial by trial and Appendix 8 in supplementary materials for details of risk of bias assessment). Omitting participant blinding from this assessment (as it is unrealistic to expect this for such and intervention) all included studies would still be at high risk of bias.

Of the 16 included studies, twelve provided data that could be included in meta-analysis. The remaining four studies (Raneri et al., 2017; Lakzadeh et al., 2016; Low et al., 2007; Schreinemachers et al., 2016) reported income, cost of intervention and dietary diversity. Results are shown study by study in Supplementary Tables 2 and 3

Primary Outcomes

Effects of home food production on xerophthalmia, night blindness and mortality in children less than 5 years old

This review found no evidence of effects of home food production on these outcomes as no trials assessed or reported them.

Effect of home food production on stunting (height-for-age) in children less than 5 years of age

Evidence of low-certainty showed that home food production may improve height-for-age in children.

Eight studies (Gelli et al., 2018; Osei et al., 2015; Olney et al., 2015; Khamhoung et al., 2000, Reinbott et al., 2016; Kuchenbecker et al., 2017; Olney et al., 2009; Marquis et al., 2017) reported on stunting and were all included in meta-analysis. Home food production increased height-for-age in children compared to control arm children in the highest quality data adjusted for clustering and other factors (mean difference MD 0.13 (z score), 95% CI 0.01 to 0.24, six studies, 5469 participants, $I^2 = 84\%$, Figure 3). This finding was supported by data adjusted for clustering only (MD 0.24 (z score), 95% CI 0.00 to 0.48, $I^2 = 41\%$), though the lowest quality data, unadjusted for clustering, did not (MD 0.03 (z score), 95% CI -0.05 to 0.12, $I^2 = 0\%$). However, sensitivity analysis using fixed-effects analysis produced differing results (data adjusted for clustering and other factors - MD 0.00 (z score), 95% CI -0.01 to 0.01).

Heterogeneity was partly explained by subgrouping by type of intervention. A positive effect was suggested in studies that combined home gardening and poultry keeping (MD 0.17 (z score), 95% CI -0.03 to 0.32, $I^2 = 86\%$), but less effect (MD 0.06 (z score), 95% CI -0.20 to 0.80, $I^2 = NA$) was suggested in studies that practiced only home gardening (difference in p value between subgroups P =0.02). Effect sizes did not differ by duration (P = 0.63) or continent (P = 0.77).

Some studies reported prevalence of stunting as well as using z-scores. Meta-analysis of prevalence data suggested that home food production reduced the prevalence of stunting in children (data adjusted for clustering, risk ratio (RR) 0.86, 95% CI 0.66 to 1.12, 206 participants, one study; unadjusted data RR 0.95, 95% CI 0.88 to 1.03, 3885 participants, four studies, $I^2 = 52\%$). Supplementary Table 4 shows all analyses for stunting.

GRADE assessment suggested low-certainty evidence that home gardening may reduce stunting in young children (downgraded once each for risk of bias and inconsistency, see Table 2).

Effect of home food production on wasting (weight-for-height) in children less than 5 years of age

Evidence of low-certainty showed that home food production may slightly improve weightfor-height in children.

Eight studies (Gelli et al., 2018; Osei et al., 2015; Olney et al., 2015; Khamhoung et al., 2000, Reinbott et al., 2016; Kuchenbecker et al., 2017; Olney et al., 2009; Marquis et al., 2017) reported wasting and were included in meta-analysis (Figure 4). Meta-analysis suggested a small benefit of home gardening on wasting (data adjusted for clustering and other factors -MD 0.05 (z score), 95% CI -0.04 to 0.14, $I^2 = 61\%$, five studies, 4510 participants), echoed in data adjusted only for clustering, but not unadjusted data. Heterogeneity was partially explained by subgrouping by continent, with greater effects in Asia (MD 0.59 z score, 95% CI 0.15 to 1.04, $I^2 = 48\%$) than Africa (MD 0.04 z score, 95% CI -0.03 to 0.11), (P for differences between subgroups = 0.021). We found no important differences between subgroups when subgrouping by duration (P = 0.22) or intervention type (P = 0.49). Sensitivity analysis (using fixed effects) supported this small beneficial effect of home gardening (MD 0.09 z score, 95% CI 0.08 to 0.10), as did the single study on prevalence of wasting in children with home gardening interventions (data adjusted for clustering, RR 0.91, 95% CI 0.44 to 1.87, 206 participants). Supplementary Table 5 shows all the analyses for

wasting.

GRADE assessment suggested low-certainty evidence that home gardening may slightly reduce wasting in young children (downgraded once each for risk of bias and imprecision, see Table 2).

Effect of home food production on underweight (weight-for-age) in children less than 5 years of age

Evidence of low certainty showed that home food production may slightly reduce underweight in children.

Seven studies (Kuchenbecker et al. 2017, Olney et al., 2015, Olney et al., 2009, Osei et al., 2015, Marquis et al., 2017, Gelli et al., 2018, Reinbott et al., 2016) were included in metaanalysis. Data adjusted for clustering and other factors suggested a small improvement in zscore for underweight (MD 0.07 z score, 95% CI -0.01 to 0.15, five studies, 4510 participants, $I^2 = 53\%$), supported by data adjusted only for clustering (MD 0.16 z score, 95% CI -0.02 to 0.34, two studies, 707 participants, $I^2 = 0\%$, unadjusted data (MD 0.03 z score, 95% CI -0.05 to 0.11, three studies, 2751 participants, $I^2 = 0\%$, Figure 5), and fixed effects sensitivity analysis (MD 0.05 z score, 95% CI 0.04 to 0.06). There was no important heterogeneity, and no important differences between subgroups by duration (P = 0.77), type of intervention (P = 0.18) or continent (P = 0.43).

Data on prevalence of underweight also showed a small benefit of home gardening though not statistically significant (data adjusted for clustering RR 0.82, 95% CI 0.57 to 1.19, one study, 206 participants; unadjusted data RR 0.95, 95% CI 0.86 to 1.05, four studies, 3888 participants, $I^2 = 25\%$). Supplementary Table 6 shows all the analyses for underweight.

GRADE assessment suggested evidence of low certainty that home food production slightly reduced underweight in children (downgraded for risk of bias and imprecision, Table 2).

Secondary Outcomes

Effects of home food production on serum retinol in children aged less than 5 years old

The effect of home food production on serum retinol was unclear. Three studies, all at high risk of bias (Faber et al., 2002; Kidala et al., 2000; Hotz et al., 2012 Uganda) reported serum retinol in children and were included in meta-analysis, but only one of these studies reported data adjusted for clustering. Home food production did not affect serum retinol (Figure 6, data adjusted for clustering and other factors MD -0.01umol/l, 95 % CI -0.06 to 0.05, one study, 413 participants; unadjusted data MD -0.07umol/l, CI -0.37 to 0.24, two studies, 367 participants, $I^2 = 92\%$). This was confirmed in fixed effects analysis (MD -0.01umol/l, 95% CI -0.06 to 0.05). With a single adjusted trial subgrouping was not possible, and prevalence data were not reported. Supplementary Table 7 shows all serum retinol analyses.

Effect of home food production on dietary diversity in children less than 5 years of age

Six studies (Gelli et al., 2018, Raneri et al., 2017., 2017; Reinbott et al., 2016; Marquis et al., 2017; Kuchenbecker et al., 2017and Olney et al. 2009) reported dietary diversity. Three studies (Gelli et al., 2018; Kuchenbecker et al., 2017 and Olney et al., 2009) could be included in meta-analysis (reported in a format that could be pooled statistically, Figure 7). For all studies, the higher the score, the higher dietary diversity in a population. Random effects meta-analysis showed that home food production increased dietary diversity in children compared to the control arm (standardized mean difference (SMD) 0.24, 95% CI 0.15 to 0.34, three studies, 2643 participants, $I^2 = 0\%$), supported by the exact same result from fixed effects analysis . Supplementary Table 8 shows all dietary diversity analyses.

Increased dietary diversity in children of intervention arms were supported by all three studies not included in meta-analysis (Raneri et al., 2017, Reinbott et al., 2016 and Marquis et al., 2017). Raneri et al., 2017showed 18% increase in minimum dietary diversity score in children in the intervention arm compared to the control arm (in the intervention arm, minimum dietary diversity increased by 0.4, P <0.01). Reinbott et al., 2016 reported mean child dietary diversity score 3.9 (SD 1.5) for the intervention arm and 3.7 (SD 1.5) for the control arm (921 children). Marquis et al., 2017 reported minimum dietary diversity of 80.2% in the intervention arm and 69.5% in the control arm (500 children, P = 0.02).

Although consumption of vitamin A-rich foods was not a pre-specified outcome of this review, post hoc we felt collecting and reporting these data may be useful. Eight studies (Kuchenbecker et al., 2017; Low et al., 2007; Hotz et al., 2012 for Mozambique; Lakzadeh et al., 2016; Hotz et al., 2012 for Uganda; Raneri et al., 2017., 2017; Reinbott et al., 2016; Marquis et al., 2016) showed that home food production increased the consumption of vitamin A-rich foods. However, 3 studies (Lakzadeh et al., 2016; Kuchenbecker et al., 2017 and Hotz et al., 2012 for Mozambique) were not statistically significant (Supplementary Table 3).

Meta-analysis and other trials suggested that home food production improved dietary diversity and vitamin-A rich food consumption in children, but the quality of this evidence is very low. Secondary outcomes were not formally assessed using GRADE.

Effect of home food production on family income

Four studies (Olney et al., 2009; Schreinemachers et al., 2016; Lakzadeh et al., 2016; Low et al., 2007) reported on income generated by home food production through sale of surplus produce but could not be pooled statistically as they lacked useable variance data. These

studies all showed that home food production can generate additional income for the household. Low et al., 2007 reported a mean revenue from home food production US\$ 3.17 ± 2.91 from orange sweet potato sales in Mozambique. Olney et al., 2009 showed that household income increased in home food production by 14.2% (P <0.05) compared to the control arm in Cambodia. Lakzadeh et al., 2016 reported a mean income of 1. 58 (P < 0.001) from home garden and fishpond in Cambodia. Schreinemachers et al., 2016 reported a mean income of -1.4 (P = 0.798) in Bangladesh (See Supplementary Table 3).

Cost of intervention

The cost of setting up a home garden was reported by two studies (Schreinemachers et al., 2016; Lakzadeh et al., 2016). Schreinemachers et al., 2016 reported that the cost of setting up a home garden in Bangladesh (project costs, women's opportunity cost and seedlings) was \$23.2 USD per annum per garden while Lakzadeh et al. (2010) reported a cost of \$220 USD for 22 months per garden and \$239 for a garden, fish pond and training per household in Cambodia (Supplementary Table 3).

Discussion

This systematic review aimed to assess effects of home food production on nutritional blindness, mortality, anthropometric measures, vitamin A status and dietary diversity in young children. No studies assessed xerophthalmia, night blindness or mortality, but evidence was found amongst the 16 included trials (2498 children) for the remaining outcomes. Evidence of low-certainty showed that introduction of home food production may slightly improve height-for-age (stunting), weight-for-height (wasting) and weight-for-age (underweight) in young children. Clinically, for a 24-month old boy, to move from a heightfor-age z-score of -2 to -1, 3.1cm in height is achieved, or 3.2cm for a girl (WHO, 2006). A boy of 24 months needs a weight gain of 1.1kg to move from a weight-for-age z-score of -2 to -1 and a girl 1.2kg. At a height of 109.5cm, 1.3kg and 1.4kg weight gain will move a boy or girl from a weight-to-height z score of -2 to -1 respectively (WHO, 2006). Limited evidence suggested no effect on children's serum retinol levels, but an increase in children's dietary diversity and household income.

This review included only studies that had an intervention and control arm. Despite this, all the studies were assessed as being of poor methodological quality. GRADE assessments were mainly downgraded due to high risk of bias of the included studies and the wide confidence intervals of the results (high levels of imprecision). Most studies were unclear on whether allocation concealment was adequate, which throws the studies open to selection bias. Most studies were unclear on whether or how participants, personnel and outcome assessors were blinded. Many of the included studies did not adjust for their clustered methodology. Overall, the evidence in this review is of low quality.

We identified seven on-going trials that will add to existing evidence and may change the findings of this review (Supplementary Table 1). This systematic review attempted to limit bias in our own methodology by adhering strictly to Cochrane methods of conducting systematic reviews. From the screening of titles to data extraction, two independent reviewers were involved, and conflicts were all resolved through discussion. Protocols were not found for most of the included studies, so it was not possible to assess outcome reporting bias. It is possible we missed some studies published in languages other than English, or published in the grey literature, such as government and charity websites.

Comparison with results of other research

A systematic review by Masset et al. (2012) included 23 clinical controlled trials and assessed the effectiveness of agricultural interventions on the nutritional status of children in low-and middle-income countries. The review found little evidence that home gardening interventions had positive effects on children's serum retinol (MD 2.4 μ g/dL, 95% CI 1.67 to 3.16). Nineteen included studies reported that home gardening improved dietary diversity, however they could not summarize across studies as the included studies measured dietary diversity in different ways. Also, the review found limited evidence that home gardening interventions increased household incomes but did not provide details on the extent of increase.

A similar systematic review by Girard et al. (2012) assessed the effectiveness of agricultural interventions on nutritional outcomes in children and women. It included 36 studies of which 32 reported on nutritional outcomes for children, finding inconsistent results for vitamin A status. All were quasi-experimental apart from one RCT. School gardening has been recently systematically reviewed by Ohly et al. (2016) using both quantitative and qualitative evidence. Ohly and colleagues included 40 studies. The quantitative evidence was of poor quality and reported that there was limited improvement in the intake of fruits and vegetables. Qualitative evidence was of a higher quality and demonstrated healthier food consumption in children with school gardening programmes.

Our systematic review was limited to higher quality trials and included more recent research. Our results are similar to those of Masset et al. (2012) and Girard et al. (2012). However, our review differed from the previous reviews regarding children's serum retinol. While they found small positive or inconsistent effects on serum retinol, our review reports that home food production did not affect serum retinol in children. Neither Girard et al. (2012) or

Masset et al. (2012) reported blindness-related outcomes. To the best of our knowledge, our systematic review is the first to investigate the impact of home food production on nutritional blindness in children, attempting to assess blindness-related outcomes. We did not find studies that reported blindness-related outcomes.

Our review suggested that home gardening including poultry production may be more effective at reducing stunting than home gardening alone. Prado et al. (2020) conducted an RCT that supplied one egg per day for six months to children aged 6 to 15 months in Malawi. They found no effect of the egg in an eye-tracking task.

Implications for Practice

This systematic review found that home food production may slightly reduce stunting, wasting and underweight in children. No study reported effects of home food production on xerophthalmia, night blindness or mortality in children, and no effects on serum retinol were found. Hence, while home food production can support anthropometric improvements, it is not clear whether it can reduce nutritional blindness in children. Better-evidenced interventions such as vitamin A supplementation programs should be adopted and expanded to children at greatest risk to prevent nutritional blindness until there is enough evidence on effects of home food production. However, introduction of home food production to improve nutritional stunting, wasting and underweight may be appropriate in areas where these are prevalent and more intensive nutritional support is not available. Home food production may be needed in low and middle income countries as 35 low and middle income countries showed a high prevalence of stunting (38.8%, 95% CI 38.6%-38.9%), wasting (12.9%, 95% CI, 12.8%-13.0%) and underweight (27.5% 95% CI 27.3% -27.6%) in 299, 353 children (Li et al., 2020).

Implications for Research

Effects of home gardening on serum retinol were inconclusive as few studies of poor quality reported on it. Large-scale high-quality trials that actively minimize selection bias, performance bias and detection bias are needed to assess effects of home food production on serum retinol as well as ophthalmic outcomes (such as night blindness, conjunctival dryness, bitot's spots), mortality, stunting, wasting and underweight in children. For example, using a common center for collection of data (rather than doing data collection at the houses of participants) and separating intervention and control geographically may help to avoid performance and detection bias. Large sample sizes, including larger numbers of small clusters, should be used to improve sample size issues, and correct adjustment for clustering effects should be routine. Adjusting for clustering prevents overestimation of statistical significance of research impact thereby avoiding biased results. One major advantage of clustering is that contamination is appropriately managed, and it is an effective way of measuring the overall effects of an intervention across a population.

Researchers should record and analyze other factors that might affect the impact of home food production on nutritional status such as deworming, environmental sanitization and potable water availability, alongside the formal results of the trial, ensuring that they are randomly distributed amongst the clusters. Even better, a factorial design could be used to assess effects of improving these factors in conjunction with home food production on nutritional blindness in children.

Choosing the right crops for the intervention (ensuring that many are rich in Vitamin A, and that they can be consumed for a large part of the year) and collection of data at the right time of the year is important when planning trials. Interventions should be culturally specific in addressing barriers and facilitators to home gardening, rallying community support, ensuring

water supplies, training on foods high in vitamin A, cooking lessons for these foods and ensuring these are locally enjoyed and accepted. Future large trials should also assess and report cost-effectiveness of home food production, and barriers and facilitators in its implementation.

Conclusion

Our review has shown that home food production may be useful in enhancing anthropometric measures in children. Home food production may, if shown to be cost-effective, be important as an integral part of other evidence-based interventions such as vitamin A supplementation in tackling stunting, wasting and underweight in children. It is likely to have a role particularly in rural areas where children are not well covered by vitamin A supplementation programs. This review may be helpful towards the global action plan launched in 2020 by the United Nations partners to urgently act towards eliminating wasting in children (WHO, 2020). Few studies, all of poor quality reported the effect of home food production on serum retinol in children. We consider the results on serum retinol inconclusive until high quality trials are carried out

Conflicts of interest

The authors state that they have no conflict of interests

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Study (reference)	Country	Duration	Setting	Study	No. of	Age of	Age of	Type of intervention
		of study		design	clusters	children	women	
Faber 2002	South	2 years	Rural	ССТ	NR	2-5	NR	Training in home gardens and nutrition education. Control arm received no
(Faber et al., 2002)	Africa					years		intervention.
Gelli 2018	Malawi	1 year	Rural	Cluster	10	6 – 72	>14	Training in agricultural practices and distribution of chicks and seedlings.
(Gelli et al., 2018)				RCT		months		Loans granted to households, cooking sessions, nutrition education. Control group was exposed to child nutrition education
Hotz 2012	Mozam-	3 years	Rural	Cluster	72	6-35	Mean	Distribution of orange sweet potato vines and nutrition education, demand
Mozambique	bique			RCT	clusters	months	age	creation. Control was exposed to no intervention
(Hotz et al., 2012)							28.9	
Hotz 2012 Uganda	Uganda	2 years	Rural	Cluster	NR	6 - 35	Mean	Distribution of orange sweet potato vines and nutrition education, demand
(Hotz et al., 2012)				RCT		months. 3	age of	creation. Control group had no intervention
						– 5 years	34.0	

Table 1. Characteristics of included studies

Khamhoung 2000	LAO	2 years	Rural	CCT		Preschool	15-45	Training on setting up home gardens and animal rearing. Control group
(Khamhoung et al.,						children	years	received no intervention
2000)								
Kidala 2000	Tanzania	2 years	Rural	ССТ	NR	6-71	NR	Training and distribution of seedlings, nutrition education, cooking
(Kidala et al., 2000)						months		sessions. Control arm received no intervention
Kuchenbecker 2017	Malawi	3 years	Rural	Cluster	24	6 – 23	Mean	Distribution of farming items, livestock and training in farming. Nutrition
(Kuchenbecker et al.,				RCT	clusters	months	age	education and cooking sessions. Control arm received only agricultural
2017)							27.2	practices with no nutrition education
Lakzadeh 2016	Cambodia	22 months	Rural	Cluster	60	<5 years	NR	Training and distribution of seedlings for home gardening. Creation of
(Lakzadeh et al., 2016; Talukder et al., 2017,				RCT	clusters			fishponds. 3 arms – home food production (HFP) plus fishpond, HFP only and control with no intervention
Verbowski et al., 2018								
Low 2007	Mozambiq	2 years	Rural	ССТ	NR	< 39	NR	Training and distribution of orange sweet potato vines, demand creation,
(Low et al., 2007)	ue					months		nutrition education. Control group was not exposed to the interventions

Marquis 2017	Ghana	1 year	rural	Cluster	16	0-32	NR	Training, distribution of seedlings, chicks and orange sweet potato vines,
(Marquis et al., 2017; Atuobi-Yeboah et al., 2016)				RCT	clusters	months		cooking sessions, nutrition education. Control group received no intervention
Olney 2009 (Olney et al., 2009)	Cambodia	19 months	Rural	RCT	NR	>5 years	NR	Training and distribution of seedlings and chicks, nutrition education. Control arm received no intervention
Olney 2015 (Olney et al., 2013; Olney et al., 2015)	Burkina Faso	2 years	Rural	Cluster RCT	NR	3 – 12.9 months	NR	Training in home garden, distributions of seedlings and chicks, nutrition education. Control arm received no intervention. Nutrition education was carried out by two groups of women – health committee and older women group.
Osei 2015 (Osei et al., 2015; Osei et al., 2017; Pries et al., 2013)	Nepal	4 years	rural	Cluster RCT	63	0 – 23 months	NR	Training in home gardening, and poultry. Nutrition education. Three arms were used- HFP, HFP plus micronutrient powder and control group that received no intervention.
Raneri 2017 (Raneri et al., 2017.)	Vietnam	1 year	Rural	Cluster RCT	NR	12 – 24 months	NR	Training in home garden, nutrition education and cooking demonstrations. Control group had no intervention.

Reinbott 2016	Cambodia	2 years	Rural	Cluster	NR	0 – 23	NR	Training in home gardening, nutrition education and giving out of
(Reinbott et al., 2016; Reinbott et al., 2018)				RCT		months		vouchers. Control arm received agricultural practices with no nutrition education
Schreinemacher 2016	Banglades	3 years	Rural	CCT	NR	Entire	NR	Training in home gardening, distribution of seedlings and orange sweet
(Schreinemacher et al.,	h					household		potato vines. Control arm received no intervention
2016)								

CCT – Controlled clinical trial (not randomized)

HFP – Home Food Production

RCT – Randomized Controlled Trial

NR - Not Reported

Table 2. Quality of evidence using GRADEpro GDT

Question: Home food production compared to control for home food production

			Certainty	y assessment			№ of pa	tients		Effect		
№ of studies	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations	home food production	control	Relative (95% CI)	Absolute (95% CI)	Certainty	Importance
Xeroph	halmia											
0	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	Critical
Night bl	indness											
0	NR	NR	NR	NR	NR	NR	NR	NR	-NR	NR	NR	Critical
Mortali	ty											
0	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	Critical
Stunting	g, z-score	e, GIV - a	adjusted for clu	stering and otl	her factors							
7	RCT	serious c	serious ^d	not serious	not serious	none	2991	2478	-	MD 0.13 higher (0.01 higher to 0.24 higher)	⊕⊕⊖⊖ LOW	IMPORTANT
Wasting	g GIV - A	djusted	for clustering a	nd other facto	rs	•			-1			
6	RCT	serious c	not serious	not serious	serious ^e	none	2498	2012	-	MD 0.05 higher (0.04 lower to 0.14 higher)	⊕⊕⊖⊖ LOW	IMPORTANT
Underw	eight GI	V - Adju	sted for cluster	ing and other t	factors		•	•	•			•
6	RCT	serious c	not serious	not serious	Serious ^e	none	2498	2012	-	MD 0.07 higher (0.01 lower to 0.15 higher)	⊕⊕⊖⊖ LOW	IMPORTANT

CI: Confidence interval; **MD:** Mean difference; **SMD:** Standardized mean difference; **RR:** Risk ratio **Explanations**

a. All the 3 studies were at high risk of bias for almost all the domains

b. Heterogeneity was high at 90%

c. All the studies had a high risk of bias for at least one domain

d. High heterogeneity at 85%e. Result was not statistically significant







Figure 2. Risk of bias summary

- represents high risk of bias, ? represents unclear risk of bias, + represents low risk of bias

Figure 3. Effects of home food production on stunting in children (z-score): Meta-analysis

assessing mean difference using the random effects model

			home food production	Control		Mean Difference	Mean Difference	Risk of Bias
Study or Subgroup	Mean Difference	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABCDEFG
1.4.1 adjusted for cluster	ing and other factor	S						
Gelli 36 - 72 mo 2018 (1)	0.05	0.05	631	617	19.5%	0.05 [-0.05, 0.15]		
Gelli 6 - 24mo 2018 (2)	0.44	0.16	155	149	8.7%	0.44 [0.13, 0.75]	Sheet Barrier	+ • • • ? ? • • • •
kuchenbecker 2017 (3)	0.17	0.12	493	466	11.9%	0.17 [-0.07, 0.41]	19 19 19 19 19 19 19 19 19 19 19 19 19 1	• ? ? ? • ? •
Marquis 2017 (4)	0.22	0.06	287	213	18.4%	0.22 [0.10, 0.34]		0200000
Olney 2015 (5)	-0.0017	0.006	801	530	22.5%	-0.00 [-0.01, 0.01]	13 4 .3	222023
Osei 2015 (6)	1.565	0.4413	100	106	1.7%	1.56 [0.70, 2.43]	2	+ • ? • ? • ? •
Reinbott 2016 (7)	-0.06	0.0714		397	17.1%	-0.06 [-0.20, 0.08]		$\bullet ? \bullet \bullet \bullet ? \bullet$
Subtotal (95% CI)			2991	2478	100.0%	0.13 [0.01, 0.24]	•	
Heterogeneity: Tau ^z = 0.02	2; Chi² = 37.26, df = 6	(P < 0.0	10001); I² = 84%					
Test for overall effect: Z = 2	2.08 (P = 0.04)							
1.4.2 adjusted for cluster	ing							
Marquis 2017	0.14	0.11	287	213	60.3%	0.14 [-0.08, 0.36]		
Osei 2015	0.39	0.1582		106	39.7%	0.39 [0.08, 0.70]		• • • • • • • • • • • •
Subtotal (95% CI)			387	319	100.0%	0.24 [-0.00, 0.48]		
Heterogeneity: Tau² = 0.01 Test for overall effect: Z = 1		(P = 0.19	l); l² = 41 %					
1.4.3 not adjusted for clus	sterina							
Khamhoung 2000	A218114	0.2551	422	102	3.0%	0.04 [-0.46, 0.54]	······································	2000920
kuchenbecker 2017		0.0714	1 (A) (C)	0.377	38.8%	0.06 [-0.08, 0.20]		
Olney 2009	-0.1	0.1225	6 83-5-5	194	13.2%	-0.10 [-0.34, 0.14]		
Olney 2015	0.05	0.0663		530	45.0%	0.05 [-0.08, 0.18]		2220929
Subtotal (95% CI)			1980	1291	100.0%	0.03 [-0.05, 0.12]	*	
Heterogeneity: Tau ² = 0.00); Chi ² = 1.39, df = 3 (P = 0.71); I ^z = 0%					
Test for overall effect: Z = 0	00090000000 V20000000 0							
								_
							-0.5 -0.25 0 0.25 0.5	

Test for subgroup differences: Chi² = 3.34, df = 2 (P = 0.19), l² = 40.1% <u>Footnotes</u>

(1) adjusted for clustering, age and sex

(2) adjusted for clustering, age and sex

(3) adjusted for sex, age, care giver education, wealth and mother's height

(4) adjusted for clustering, child age, sex, baseline values, maternal and household factors

(5) adjusted for clustering, maternal height, sex, age and maternal age

(6) adjusted for clustering, age, sex, maternal characteristics, household wealth

(7) adjusted for clustering, age of child, household wealth, maternal education

Risk of bias legend

favours control favours HFP

(A) Random sequence generation (selection bias)

(B) Allocation concealment (selection bias)

(C) Blinding of participants and personnel (performance...

(D) Blinding of outcome assessment (detection bias)

(E) Incomplete outcome data (attrition bias)

(F) Selective reporting (reporting bias)

Figure 4. Effects of home food production on wasting in children (z-score): Meta-analysis

assessing mean difference using the random effects model

		hom	e food production	Control		Mean Difference	Mean Difference	Risk of Bias
Study or Subgroup	Mean Difference	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABCDEFG
1.5.1 Adjusted for clusteri	ng and other factor	S					50 C	
Gelli 36 - 72 mo 2018 (1)	-0.04	0.07	631	617	18.9%	-0.04 [-0.18, 0.10]		
Gelli 6 - 24mo 2018 (2)	-0.13	0.15	155	149	7.2%	-0.13 [-0.42, 0.16]		
Marquis 2017 (3)	0.07	0.08	287	213	16.6%	0.07 [-0.09, 0.23]	3 1 1 1 1 1 1	
Diney 2015 (4)	0.0932	0.0066	801	530	34.0%	0.09 [0.08, 0.11]	-	222020
Dsei 2015	0.595	0.227	100	106	3.6%	0.59 [0.15, 1.04]		
Reinbott 2016 (5)	0	0.0663	524		19.8%	0.00 [-0.13, 0.13]		
Subtotal (95% CI)			2498	2012	100.0%	0.05 [-0.04, 0.14]	•	
Heterogeneity: Tau ² = 0.01;	Chi ² = 12.67, df = 5	(P = 0.03); I ²	= 61%					
Fest for overall effect: Z = 1	.05 (P = 0.29)							
1.5.2 adjusted for clusteri	ng							
Marquis 2017	0.04	0.13	287	213	56.4%	0.04 [-0.21, 0.29]	· · · · · · · · · · · · · · · · · · ·	
Dsei 2015	0.09	0.148	100		43.6%	0.09 [-0.20, 0.38]	105	• ? • ? • ? •
Subtotal (95% CI)			387	319	100.0%	0.06 [-0.13, 0.25]		
Heterogeneity: Tau ² = 0.00;	Chi ² = 0.06, df = 1 (P = 0.80); I ^z =	0%					
Fest for overall effect: Z = 0	.63 (P = 0.53)							
1.5.3 Unadjusted for clust	ering							
<hamhoung 2000<="" td=""><td>0.1</td><td>0.0867</td><td>422</td><td>102</td><td>18.5%</td><td>0.10 [-0.07, 0.27]</td><td></td><td>2000020</td></hamhoung>	0.1	0.0867	422	102	18.5%	0.10 [-0.07, 0.27]		2000020
kuchenbecker 2017	0.05	0.0612	493	465	35.8%	0.05 [-0.07, 0.17]		• • • • • • • •
Olney 2009	-0.1	0.0969	264	193	14.9%	-0.10 [-0.29, 0.09]		
Diney 2015	-0.03	0.0663	801	530	30.8%	-0.03 [-0.16, 0.10]		22203
Subtotal (95% CI)			1980	1290	100.0%	0.01 [-0.06, 0.09]	+	
Heterogeneity: Tau ² = 0.00;	Chi ² = 3.15, df = 3 ($P = 0.37$; $I^2 =$	5%					
Test for overall effect: Z = 0		10.55						

Test for subgroup differences: Chi² = 0.48, df = 2 (P = 0.79), I² = 0%

Footnotes (1) adjusted for clustering, age, sex and clustering (2) adjusted for clustering, age, sex and clustering

(3) adjusted for clustering, child age, sex, baseline values, maternal and household factors
 (4) adjusted for clustering, maternal height, sex, age and maternal age
 (5) adjusted for clustering, age, household wealth and maternal education

-0.5 -0.25 0 0.25 0.5 favours control favours HFP

Risk of bias legend

(A) Random sequence generation (selection bias)

(B) Allocation concealment (selection bias)

(C) Blinding of participants and personnel (performance... (D) Blinding of outcome assessment (detection bias)

(E) Incomplete outcome data (attrition bias)

(F) Selective reporting (reporting bias)

Figure 5. Effects of home food production on underweight in children (z-score): Meta-analysis

assessing mean difference using the random effects model

		ho	me food production	Control		Mean Difference	Mean Difference	Risk of Bias
Study or Subgroup	Mean Difference	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABCDEFG
1.6.1 Adjusted for cluster	ing and other factor	S					5 Q	
Gelli 36 - 72 mo 2018 (1)	0.05	0.05	631	617	22.6%	0.05 [-0.05, 0.15]		
Gelli 6 - 24mo 2018 (2)	-0.02	0.14	155	149	6.8%	-0.02 [-0.29, 0.25]		
Marquis 2017 (3)	0.15	0.07	287	213	17.1%	0.15 [0.01, 0.29]		
Olney 2015 (4)	0.0525	0.0059	801	530	33.8%	0.05 [0.04, 0.06]	-	??? 🔴 🔁 ? 🧲
Osei 2015	1.055	0.3138	100	106	1.6%	1.05 [0.44, 1.67]	80	→ ●?●?●?●
Reinbott 2016 (5)	-0.02	0.0663	524	397	18.0%	-0.02 [-0.15, 0.11]		
Subtotal (95% CI)			2498	2012	100.0%	0.07 [-0.01, 0.15]	•	
Heterogeneity: Tau² = 0.00		(P = 0.02);	I ² = 63%					
Fest for overall effect: Z = 1	1.63 (P = 0.10)							
1.6.2 adjusted for clusteri	ing							
Marquis 2017	0.09	0.12	287	213	58.6%	0.09 [-0.15, 0.33]		
Osei 2015	0.27	0.1429	101	106	41.4%	0.27 [-0.01, 0.55]		• ? • ? • ? •
Subtotal (95% CI)			388	319	100.0%	0.16 [-0.02, 0.34]	-	
Heterogeneity: Tau ² = 0.00); Chi ^z = 0.93, df = 1 ((P = 0.33); I ²	= 0%					
Fest for overall effect: Z = 1	1.79 (P = 0.07)							
1.6.3 Unadjusted for clust	tering							
kuchenbecker 2017	0.07	0.0663	493	466	37.9%	0.07 [-0.06, 0.20]		• ? ? ? • ? •
Olney 2009	0	0.0969	266	195	17.7%	0.00 [-0.19, 0.19]	+	
Olney 2015	0	0.0612	801	530	44.4%	0.00 [-0.12, 0.12]		????
Subtotal (95% CI)			1560	1191	100.0%	0.03 [-0.05, 0.11]	+	
Heterogeneity: Tau ² = 0.00); Chi ² = 0.69, df = 2 (P = 0.71); I ²	= 0%					
Test for overall effect: Z = 0	0.65 (P = 0.52)	10.05						
	50 SA							

Test for subgroup differences: Chi² = 1.98, df = 2 (P = 0.37), l² = 0% <u>Footnotes</u>

(1) adjusted for clustering, age and sex

(2) adjusted for clustering, age and sex

(3) adjusted for clustering, child age, sex, baseline values, maternal and household factors

(4) adjusted for clustering, maternal height, sex, age and maternal age

(5) adjusted for clustering, age, household wealth and maternal education

Risk of bias legend

-0.5 -0.25 0 0.25 0.5 favours control favours HFP

(A) Random sequence generation (selection bias)

(B) Allocation concealment (selection bias)

(C) Blinding of participants and personnel (performance...

(D) Blinding of outcome assessment (detection bias)

(E) Incomplete outcome data (attrition bias)

(F) Selective reporting (reporting bias)

Figure 6. Forest plot showing serum retinol (µmol/l) in children using random effects meta-

analysis

		home f	ood production	Control		Mean Difference	Mean Difference	Risk of Bias
Study or Subgroup	Mean Difference	SE	Total	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% CI	ABCDEFG
1.14.1 adjusted for c	clustering and other	factors						
Hotz 2012 Uganda Subtotal (95% CI)	-0.005	0.03	205 205		100.0% 100.0%	-0.01 [-0.06, 0.05] -0.01 [-0.06, 0.05]		229929
Heterogeneity: Not a	pplicable							
Test for overall effect	: Z = 0.17 (P = 0.87)							
1.14.2 not adjusted I	for clustering							
Faber 2002	0.08	0.0255	110	111	53.1%	0.08 [0.03, 0.13]	-	• ? ? ? ? ? •
Kidala 2000	-0.23	0.0816	75	71	46.9%	-0.23 [-0.39, -0.07]		0000020
Subtotal (95% CI)			185	182	100.0%	-0.07 [-0.37, 0.24]		
Heterogeneity: Tau ² =	= 0.04; Chi ² = 13.15,	df = 1 (P = 0.000	3); ² = 92%					
Test for overall effect	: Z = 0.42 (P = 0.67)							
							40 1/2 1/2 1/2	
							-0.5 -0.25 0 0.25 0.5	_
							Favours control Favours HFP	
Test for subgroup dif	fferences: Chi² = 0.1	5, df = 1 (P = 0.70)), I² = 0%					
Dick of bigs lagand								

Risk of bias legend

(A) Random sequence generation (selection bias)

(B) Allocation concealment (selection bias)

(C) Blinding of participants and personnel (performance bias)

(D) Blinding of outcome assessment (detection bias)

(E) Incomplete outcome data (attrition bias)

(F) Selective reporting (reporting bias)

Figure 7. Effects of home food production on dietary diversity in children (z-score): Meta-

analysis assessing mean difference using the random effects model

Study or Subgroup	Std. Mean Difference		home food production Total		Weight	Std. Mean Difference IV, Random, 95% Cl	Std. Mean Difference IV, Random, 95% Cl	Risk of Bias A B C D E F G
.12.1 not adjusted for clus	stering							
elli 36 - 72 mo 2018 (1)	0.2298	0.0577	606	604	65.4%	0.23 [0.12, 0.34]		
uchenbecker 2017 (2)	0.39	0.15	493	466	9.7%	0.39 [0.10, 0.68]		
Diney 2009 (3) Subtotal (95% CI)	0.2229	0.0935	277 1376	197 1267	24.9% 100.0%	0.22 [0.04, 0.41] 0.24 [0.15, 0.34]	•	•?????
leterogeneity: Tau ² = 0.00; 'est for overall effect: Z = 5.3	20~622121).59); I² =		1207	100.0%	0.24 [0.15, 0.34]	-	

Test for subgroup differences: Not applicable

Footnotes

(1) A 24-hour recall method was used where energy, protein, vitamin A and other micronutrients were assessed (2) dietary diversity was calculated using 10 food groups

(3) A 3-day recall was used and dietary diversity was assessed across 6 food groups.

Risk of bias legend

-0.5 -0.25 0 0.25 0.5 Favours control Favours HFP

(A) Random sequence generation (selection bias)

(B) Allocation concealment (selection bias)

0.25 0.5

(C) Blinding of participants and personnel (performance...

(D) Blinding of outcome assessment (detection bias)

(E) Incomplete outcome data (attrition bias) (F) Selective reporting (reporting bias)