THE GENETIC ARCHITECTURE OF SEXUAL DIMORPHISM IN THE MOSS CERATODON PURPUREUS

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ABSTRACT

A central problem in evolutionary biology is to identify the forces that maintain genetic variation for fitness in natural populations. Sexual antagonism, in which selection favors different variants in males and females, can slow the transit of a polymorphism through a population or can actively maintain fitness variation. The amount of sexually antagonistic variation to be expected depends in part on the genetic architecture of sexual dimorphism, about which we know relatively little. Here, we used a multivariate quantitative genetic approach to examine the genetic architecture of sexual dimorphism in a scent-based fertilization syndrome of the moss *Ceratodon purpureus.* We found sexual dimorphism in numerous traits, consistent with a history of sexually antagonistic selection. The cross-sex genetic correlations (r_{mf}) were generally heterogeneous with many values indistinguishable from zero, which typically suggests that genetic constraints do not limit the response to sexually antagonistic selection. However, we detected no differentiation between the female- and male-specific trait (co)variance matrices ($\mathbf{G_f}$ and $\mathbf{G_m}$, respectively), meaning the evolution of sexual dimorphism may be constrained. The cross-sex cross-trait covariance matrix **B** contained both symmetric and asymmetric elements, indicating that the response to sexually antagonistic or sexually concordant selection, and the constraint to sexual dimorphism, is highly dependent on the traits experiencing selection. The patterns of genetic variances and covariances among these fitness components is consistent with partly sex-specific genetic architectures having evolved in order to partially resolve multivariate genetic constraints (i.e. sexual conflict), enabling the sexes to evolve toward their sex-specific multivariate trait optima.

INTRODUCTION

Males and females achieve fitness through different strategies [1–3], which can drive the evolution of sexual dimorphism [4,5]. The ubiquity of sexual dimorphism suggests that selection frequently favors different trait optima in males and females. Sexual conflict occurs when an allelic substitution that increases fitness in one sex decreases fitness in the other, and thus both sexes are prevented from reaching their respective fitness optimum [6]. Theory and empirical evidence show that opposing selection in males and females can maintain genetic variation for fitness [7–15]. However, whether sexual conflict in a population is evolutionarily transient or persistent will depend on both the nature of sex-specific selection and the nature of sex-specific genetic architecture for traits [16–20] the latter of which remains poorly understood, especially in non-model organisms.

The simplest means to evaluate the constraint imposed by a shared underlying genetic architecture for homologous traits between the sexes is to measure the cross-sex genetic correlation (r_{mf}) [6]. A strongly positive r_{mf} for a trait will cause selection in one sex to generate a correlated response in the other sex [4,21] precluding the evolution of sexual dimorphism. Poissant et al. [22] found that half of the estimates of r_{mf} in 114 studies were above ~0.8, indicating that sexual dimorphism may often be constrained by traits having shared genetic architecture in males and females. Additional evidence for constraint on the evolution of sexual dimorphism is provided by studies identifying opposing selection gradients on correlated traits [22,23]. The resolution of sexual conflict can occur by the evolution of sex linkage or various forms of sex-biased gene expression (sex-specific genetic modifiers, and genomic imprinting) [4,24–26] and allows differential response to selection in males and females.

Single trait analyses, however, fail to account for covariances among traits within and between the sexes, which are important for predicting the response to selection [21]. The multivariate constraint to sexual dimorphism is captured by the sex-specific genetic variance-covariance matrix ($\mathbf{G_{mf}}$), which represents a more complete framework for studying genetic

architecture [18,27,28]. $\mathbf{G_{mf}}$ consists of the female- and male-specific submatrices $\mathbf{G_f}$ and $\mathbf{G_m}$, respectively, as well as the cross-sex cross-trait covariance matrix, **B** (and its transpose, $\mathbf{B^T}$):

$$\mathbf{G_{mf}} = \begin{bmatrix} \mathbf{G}_f & \mathbf{B} \\ \mathbf{B}^T & \mathbf{G}_m \end{bmatrix}. \quad \text{Eq. 1}$$

The diagonals of $\mathbf{G_m}$ and $\mathbf{G_f}$ represent the genetic variances of the traits in males and females, respectively, and the off-diagonals within $\mathbf{G_m}$ or $\mathbf{G_f}$ are the sex-specific genetic covariances between pairs of traits. The within-trait cross-sex covariances along the diagonal of the **B** matrix can be standardized into estimates of r_{mf}, while the off-diagonal elements of **B** represent the cross-sex cross-trait covariances – i.e. covariances between a trait in one sex and a different trait in the opposite sex. While $\mathbf{G_m}$ and $\mathbf{G_f}$ are symmetric matrices, **B** is a square matrix that may not be symmetrical (i.e., **B** need not equal $\mathbf{B^T}$). Asymmetries in **B** may play an important role in the evolution of sexual dimorphism, although the prevalence of such asymmetry is unknown outside of a few model systems [20,29].

The moss *Ceratodon purpureus* is an emerging model for studying sex-specific genetic architecture. Nearly 60% of moss species have separate males and females, and sexual dimorphism is common, most notably in the production of volatile organic compounds (VOCs) [30]. *Ceratodon purpureus* females produce a wider variety and greater quantity of VOCs than males. In choice experiments with *C. purpureus*, microarthropods, such as mites and springtails, were more attracted to female than male moss VOCs [29]. Furthermore, co-cultivating mosses with microarthropods increases moss fertilization success by ~5X [31]. These observations suggest that mosses and microarthropods are engaged in scent-based fertilization analogous to pollinator mutualisms in flowering plants. An increase in VOC production may attract more sperm-dispersing arthropods, enhancing both fertilization and the opportunity for mate choice

[32]. In males, however, VOC production may expend resources that could be allocated to other fitness components (e.g., sperm production). Thus, the evolution of VOC production toward sex-specific fitness optima could conceivably be limited by genetic covariances between traits, sexes and trait/sex combinations.

The moss system has several technical features that make it an excellent model for sex-specific quantitative genetic analyses. The dimorphic part of the life cycle is haploid, meaning there is no dominance component of genetic variation in dimorphic traits. Sex in this system is determined at meiosis, by the segregation of the U and V sex chromosomes (as opposed to XY/ZW systems, where sex is determined at fertilization). The diploid sporophyte is always heterozygous (i.e., UV). This is because only the haploid male gametophytes make sperm, and only the female gametophytes make eggs – each chromosome is transmitted through only one sex. At meiosis, spores inheriting a U develop into female haploid gametophytes, while spores inheriting a V are males [31]. Thus, each sex contains a non-recombining sex-limited chromosome, meaning that the various asymmetries associated with the sex chromosome content in XY or ZW systems are absent [33]. Finally, the gametophytes are clonally replicable, which enables large sample sizes and limits environmental variation, increasing statistical power to estimate genetic (co)variances.

Here, we take advantage of these features to study the genetic architecture of multivariate sexual dimorphism in a natural population of the moss *C. purpureus*. We estimate $\mathbf{G}_{mf}$ and explicitly compare the male and female variance-covariance matrices, test for asymmetry in B, and compare the results of single-trait and multi-trait analyses. The cross-sex correlations were heterogeneous across traits and mostly indistinguishable from zero, suggesting that the evolution of sexual dimorphism is relatively unconstrained. We detected no differences between the female and male (co)variance matrices ($\mathbf{G}_f$ and $\mathbf{G}_m$), suggesting the sexes are likely to exhibit a similar response to selection. However, this in combination with asymmetry in the **B** matrix indicates that even sexually concordant selection could generate

sexual dimorphism. Nevertheless, **B** also contained symmetric components, suggesting possible ongoing sexual conflict in the form of lasting, unresolved constraints to the evolution of further sexual dimorphism.

Materials and Methods

Haploid sibling family cultivation

To generate a genetically diverse sample of haplotypes to estimate the phenotypic and genetic variation in *C. purpureus*, we generated axenic cultures of 45 haploid sibling families each consisting of a minimum of 3 male and 3 female siblings [34]. These families were generated from 45 sporophytes collected in Portland, OR, with each sporophyte representing a single family. This design is analogous to genotyping the sperm from a single male in an XY system which allows us to compare the underlying genetic architecture of male and female traits within a family and understand sex specific differences.

To establish axenic lines from field-collected plants, we surface-sterilized operculate sporophytes and created spore solutions following published protocols [35,36]. We plated 100 µL of the spore suspension on BCD media with 0.5 mM ammonium tartrate [37]. We germinated spores under fluorescent lights (18 hours dark and 6 hours light) and isolated single haplotypes. We confirmed sex following Norrell et al. [38] and by observing sex structures.

Collection of growth, development, morphology, and physiology traits

We grew a total of five replicates from 345 genotypes. We grew two replicates in a greenhouse in Portland, OR. From these plants we collected volatiles at peak sex expression, as this is when the moss was observed to be most fragrant. Following volatile collection (see below), we calculated a dry weight, analyzed leaf measurements using automated methods in ImageJ, and dissected tissue to confirm the presence of sex structures, measure reproductive effort, and eliminate non-sex expressing profiles.

147 We used the remaining three replicates in a common growth chamber experiment to
148 survey variation in growth and development. We grew each genotype on BCDA media, following
149 Burtscher et al.[39]. Starting on day 0 and every 7 days after for 21 days, we collected
150 measurements of juvenile growth (protonema) and development, including area, perimeter, and
151 circularity (a measure of how much the growth pattern deviated from a perfect circle (C;
152 Supplemental methods Eq. S1)). Protonemal growth patterns in which the measured perimeter
153 matched the estimated perimeter (assuming that the measured area was a perfect circle) return
154 C = 1, while growth patterns with larger measured perimeters (e.g., more star-shaped) return
155 values C < 1. Plants with circularity near 1 are largely comprised of chloronema (less mature
156 cell type). Having a larger perimeter relative to area (C < 1) suggests more mature, longer
157 celled caulonema, and indicates faster maturation. Throughout this manuscript, we refer to
158 perimeter and circularity of protonemal tissue after 21 days of growth as “juvenile growth” and
159 “juvenile growth form”, respectively. We also observed the accumulation of mature leafy
160 gametophores after 21 days, recording the total number of gametophores present. We refer to
161 the accumulation of gametophores as “mature tissue”.
162
163 ***Collection of volatile organic compounds (VOCs)***
164 We sampled VOC emissions over 9 consecutive days using a proton transfer reaction
165 time of flight mass spectrometer (PTR-TOF-MS 1000, Ionicon), incorporating a custom
166 designed sampling apparatus with hydronium (H_3O+) as the primary reagent ion (ESM1 Figure
167 S1). Prior to VOC collection, we dark-adapted replicates for 12 hours and measured chlorophyll
168 fluorescence (Opti-Sciences OP5+,Hundson, New Hampshire) to assess overall plant health
169 and remove stressed plants from the study which could lead to outliers in VOC profiles. For
170 each replicate we carefully extracted 200 mg (wet weight) of mature gametophore tissue,
171 removing remnants of soil, BCDA media, and other contaminants. We placed the plant tissue in
172 5 ml vials with distilled water to avoid dehydrating the plant during static head space

accumulation. We placed all sample and blank cuvettes under an LED light source at 1000 PAR for two hours at 35°C. All 75 masses we report are protonated species; however, we represent volatile production as the number of different masses produced ("total masses") and total concentration of overall volatile production ("total concentration").

Estimating the Genetic (Co)variance Matrix

We used a multivariate framework to estimate the extent to which the shared genome between males and females imposes a constraint on the evolution of sexual dimorphism. All of these analyses involve analyzing a fitted $\mathbf{G_{mf}}$. We fit the genetic (co)variance matrix, $\mathbf{G_{mf}}$, as a random effect in a general linear mixed-effects model (GLMM) using Bayesian Markov chain Monte Carlo (MCMC) simulations in the package 'MCMCglmm' (v. 2.29, [40]). We fit two models to estimate $\mathbf{G_{mf}}$: one for growth and development traits and another representing morphology and physiology traits. Our model for growth and development traits included juvenile growth, juvenile growth form, and mature tissue, while the model for morphology and physiology traits included total masses produced, total concentration across all masses, relative reproductive effort, and leaf length. We fit two models because traits were collected on plants grown in different environments (growth chamber vs greenhouse) and at different stages. Thus, the categories of traits are arbitrary and titles for each model are simply for convenience. All traits in both models were zero-centered and variance-standardized across sexes. To account for sex specific reproductive strategies, reproductive effort was first divided by the sex-specific means (i.e., transformed to relative reproductive effort) and then zero centered and variance standardized across the sexes. Total concentration was calculated by first dividing each of the 75 detected masses by their respective means, summing the concentrations for each observation, and log transforming this sum. We used MCMCglmm()'s 'trait' function to identify our multivariate list of traits in the response variable as a fixed effect (trait), which we interacted

with the fixed effect of ‘sex’ (trait:sex) to estimate the degree of sexual dimorphism for each trait, making the full GLMM:

$$y = \text{trait-1} + \text{trait:sex} + \mathbf{G_{mf}} + \text{sampleID} + q + e \quad \text{(Eq. 1)}$$

where y is a phenotypic vector of the traits, trait-1 indicates a model fit without an intercept, $\mathbf{G_{mf}}$ was estimated over the 45 haploid sibling families ('famid’), sampleID is the random effect of clonal replicate, q is an additional random effect (see below), and *e* is the unexplained residual variance (a Gaussian error structure was assumed for all traits). The best fitting model (as inferred by DIC comparisons; see below) for growth and development was a 3-trait (6×6 $\mathbf{G_{mf}}$) where q was ‘plate’, while the best fitting model for morphology and physiology was a 4-trait (8×8 $\mathbf{G_{mf}}$) where q was ‘date’. We modeled (co)variances using the following random effects structure of MCMCglmm: random = ~us(trait:sex):famid. Residual covariances were fixed to zero (rcov = ~idh(trait:sex):units), as male and female measures were made on separate individuals.

We used parameter expanded priors (as in Grieshop et al. [41]) for the growth and development model and inverse-Gamma priors (as in Puentes et al. [42]) for the morphology and physiology model. To determine the robustness of the posterior distribution to the prior [43,44] we compared models to other priors. The joint posterior distribution was estimated from 1,000,000 MCMC iterations after a burn-in period of 5,000 iterations, and every 1,000th posterior estimate was stored – providing 1,000 uncorrelated posterior estimates for downstream $\mathbf{G_{mf}}$ analyses. Model convergence was assessed using Gelman and Rubin diagnostics [45] and through visual inspection.

Because variance estimates of **G** matrices are bounded by zero, we evaluated whether (sex-specific) genetic variances were significantly different from zero via univariate model comparisons. All univariate models were fit using the inverse-Gamma priors while keeping all else equal to the respective multivariate models. Sex-specific genetic variance was detected as a delta DIC of 2 or more [46] between models with and without the "sex" term in the random effect of $\mathbf{G_{mf}}$ (making it simply **G**), and genetic variance was detected in the same way by comparing models with and without **G** (ESM2 table S1). We conducted all statistical analyses using R (version 4.0.2; R Development Core Team 2020).

Descriptive statistics

Sex specific genetic variances, intersexual genetic correlations (r_{mf}), and sexual dimorphisms for each trait were estimated directly by our MCMC model. Male and female genetic variances were estimated on the diagonal of the two sex-specific sub matrices $\mathbf{G_f}$ and $\mathbf{G_m}$ – we report the highest posterior density (HPD) mean estimates with upper and lower 95% HPD intervals as credibility intervals (CIs) in table 1. The cross-sex genetic correlations for traits, r_{mf}, were estimated along the diagonal of the *correlation* matrix for **B** (i.e., the standardized covariances, which are estimated directly by MCMCglmm()) – we report the HPD mode r_{mf} estimates with upper and lower 95% CIs (table 1). If $r_{mf} = 1$, it means that selection acting to increase a trait value in one sex would cause a correlated response of that same trait in the opposite sex – i.e. response to selection would be constrained. Consequently, an r_{mf} of zero would enable that trait to respond to sex-specific selection with no effect in the other sex. Lastly, we report the sign (male – female) and magnitude of sexual dimorphism for each trait as the HPD means and Cis for the estimated fixed effect trait:sex, with p-values provided by MCMCglmm() (table 1).

Similarity between $\mathbf{G_f}$ and $\mathbf{G_m}$

To compare the size, shape, and orientation of $\mathbf{G_f}$ and $\mathbf{G_m}$, we calculate Hansen's difference *d* [19] and a simplified version of the eigentensor comparison [47,48]. Hansen's *d* estimates the average distance between endpoints of response vectors generated from random selection gradients on the $\mathbf{G_f}$ and $\mathbf{G_m}$ matrices [19], similar to a random skewers method [49]. An eigentensor analysis [50,51] comparing two symmetric matrices reduces to a simple difference between the matrices. Thus, we obtained an estimate of the difference between $\mathbf{G_f}$ and $\mathbf{G_m}$ by taking the difference between the 1,000 paired posterior estimates of $\mathbf{G_f}$ and $\mathbf{G_m}$ and calculating the trace (sum of the eigenvalues) of this difference matrix. We report the HPD mode and 95% CIs of that trace. A test of the significance of this difference was obtained by comparison to that of a null distribution, which was generated by randomly swapping the sex labels of the 1,000 paired $\mathbf{G_m}$ and $\mathbf{G_f}$ estimates. With the mode of these null estimates being very near zero and the true estimate being positive, the two-tailed p-value is simply the proportion the 1,000 posterior estimates of the true difference that were < their respective null estimates of the difference, times two [41]. The eigentensor comparison of $\mathbf{G_f}$ and $\mathbf{G_m}$ provided qualitatively similar results (see ESM3 figure S2).

Symmetry of B

Asymmetry in the **B** matrix indicates differences in the underlying genetic architecture for traits between the males and females [50,51]. For example, an off-diagonal element of **B** with a covariance of 1 between trait *i* in males and trait *j* in females would suggest that selection on trait *i* in males would cause a correlated response to trait *j* in females. Asymmetry in **B** means that selection on trait "*i*" in females will produce a correlated response on trait "*j*" in males, but that correlated response differs if the sexes are reversed – i.e., selection on trait *i* in males produces a different correlated response in females. Thus, the relative proportion of **B** that is symmetric versus asymmetric reveals the relative magnitude of cross-sex cross-trait pleiotropic

271 constraints versus sex-specific genetic architecture, respectively. Thus, we partitioned **B** into its
272 symmetric and asymmetric (or skew symmetric) components using matrix decomposition
273 [29,52]. Any square matrix - **A** (e.g. **B**) - is the summation of the two components **S** and **N**:
274

$$\mathbf{A} = \mathbf{S} + \mathbf{N} \quad \text{(Eq. 3)}$$

275 the symmetric and asymmetric components, respectively, where $\mathbf{S} = \frac{1}{2}(\mathbf{A} + \mathbf{A}^T)$ and $\mathbf{N} = \frac{1}{2}(\mathbf{A} - \mathbf{A}^T)$. The proportions of **B** that are symmetric and asymmetric are given by the ratio of the sums
277 of squares of those components to that of the total, **B** [51,52]. We report the HPD mode and
278 95% CIs for these proportions by resampling them from the 1,000 stored posterior estimates of
279 **B**.
280

281 **Antagonistic and concordant genetic variation**

282 To evaluate the relative proportion of genetic variation in this population that would
283 respond to sexually concordant versus sexually antagonistic selection, we estimated the matrix
284 $\mathbf{G}_{ca}$, following Sztepanacz and Houle [52]. The submatrices of $\mathbf{G}_{ca}$, $\mathbf{G}_{a}$ and $\mathbf{G}_{c}$ predict the
285 response of the sex difference in trait values to sexually antagonistic selection, and the
286 response of trait means to sexually concordant selection, respectively. We projected $\mathbf{G}_{mf}$ onto a
287 set of arbitrary orthonormal vectors ($\mathbf{S}_m$) that spanned the concordant and antagonistic
288 subspaces of $\mathbf{G}_{mf}$. If an n-trait $\mathbf{G}_{mf}$ has 2n dimensionality (e.g. 8 in the case of the 4-trait
289 morphology and physiology matrix), then $\mathbf{S}_m$ was constructed by first taking the set of n
290 eigenvectors that span the space of an n-dimensional identity matrix, dividing them (arbitrarily)
291 by the square root of two (giving $\mathbf{E}_m$), and arranging them into the following 2n-dimentional
292 matrix: $\mathbf{S}_m = \begin{bmatrix} \mathbf{E}_m & \mathbf{E}_m \\ \mathbf{E}_m & -\mathbf{E}_m \end{bmatrix}$. The unit-length vectors of the first n columns of $\mathbf{S}_m$ therefore span the

sexually concordant subspace of $\mathbf{G_{mf}}$ and the unit-length vectors of the second n columns of $\mathbf{S}_m$ span the sexually antagonistic subspace of $\mathbf{G_{mf}}$ [52]. $\mathbf{G_{mf}}$ was projected onto this space:

$$\mathbf{G_{ca}} = \mathbf{S}_m^{\mathsf{T}}\, \mathbf{G_{mf}}\, \mathbf{S}_m, \qquad \text{(Eq. 4)}$$

where the upper-left and bottom-right n-dimensional submatrices of $\mathbf{G_{ca}}$ are covariance matrices that represent the sexually concordant ($\mathbf{G_c}$) and sexually antagonistic ($\mathbf{G_a}$) subspaces of $\mathbf{G_{mf}}$, respectively [52]. The proportion of $\mathbf{G_{mf}}$ that is sexually concordant and sexually antagonistic is therefore given by the ratio of the trace of $\mathbf{G_c}$ to $\mathbf{G_{mf}}$ and $\mathbf{G_a}$ to $\mathbf{G_{mf}}$, respectively [52]. Again, we report the HPD mode and 95% CIs for these overall proportions, as well as for each eigenvector of $\mathbf{G_{mf}}$, $\mathbf{G_c}$ and $\mathbf{G_a}$, by resampling the 1,000 stored posterior estimates of $\mathbf{G_{mf}}$.

RESULTS

Sex-specific genetic variances, r_{mf}, and sexual dimorphism

We found that leaf length and total masses were sexually dimorphic in our multivariate models. The sign (male - female) and magnitude of sexual dimorphism for each trait are reported as the HPD means and CIs estimated by the trait:sex fixed effect (table 1). We identified non-zero genetic variance in all traits, and non-zero sex-specific genetic variance in all traits except leaf length (ESM2 table S1). Male and female genetic variances were estimated on the diagonal of the two sex-specific sub matrices $\mathbf{G_f}$ and $\mathbf{G_m}$ – we report the HPD mean estimates and 95% CIs in table 1. The magnitude of sex-specific genetic variances ranged from 0.1 to 0.2 in growth and development and 0.001 to 0.1 in morphology and physiology (table 1). Many of our estimated genetic covariances were strong but accompanied by large uncertainties (ESM4 figure S4, S6) which is not uncommon [40]. Juvenile growth form and leaf length had positive r_{mf} estimates with CIs that did not include zero (table 1).

Comparing $\mathbf{G_m}$ and $\mathbf{G_f}$

We used two methods to assess the overall similarity between the male and female (co)variance sub-matrices $\mathbf{G_m}$ and $\mathbf{G_f}$. Hansen's difference *d* indicated that there were broadly no differences between $\mathbf{G_m}$ and $\mathbf{G_f}$ in terms of their multidimensional size, shape or orientation for growth and development traits (*d* = 0.094, CIs: -0.043, 0.228) or morphology and physiology traits (*d* = 0.062, CIs: -0.005, 0.129) (figure 1, table 2). The simplified eigentensor analysis (as well as the formal version, ESM3 figure S2) showed that $\mathbf{G_m}$ and $\mathbf{G_f}$ were similar for both growth and development traits (difference = -0.173, CIs: -0.544, 0.121, p = 0.284) and morphology and physiology traits (difference = -0.073, CIs: -0.269, 0.091, p = 0.24) (figure 1, table 2).

Analyzing B

We estimated symmetry and asymmetry in the **B** matrix by comparing the off-diagonal elements. Across growth and development traits, the proportion of the **B** matrix that was asymmetric was 0.112 (CIs: 0.002, 0.448) and the proportion that was symmetric was 0.884 (CIs: 0.552, 0.998) (figure 1A, table 2, ESM4 figure S3,S4). Across morphology and physiology measurements, the proportion of the **B** matrix that was asymmetric was 0.312 (CIs: 0.064, 0.513) and the proportion that was symmetric was 0.688 (CIs: 0.487, 0.936) (figure 1B, table 2, ESM4 figure S5,S6).

Concordant and antagonistic subspace of $\mathbf{G_{mf}}$

For growth and development traits, proportionally 0.367 (CIs: 0.248, 0.476) of the total genetic variances laid within the antagonistic subspace while proportionally 0.633 (CIs: 0.524, 0.752) of the total genetic variances laid within the concordant subspace (table 2). For morphology and physiology traits, 0.241 (CIs: 0.121, 0.466) of the total genetic variances laid within the antagonistic subspace while 0.759 (CIs: 0.534, 0.879) laid within concordance subspace (table 2).

We plot the genetic variances for the eigenvectors of the concordant ($\mathbf{G_C}$) and antagonistic ($\mathbf{G_A}$) subspaces alongside that of $\mathbf{G_{mf}}$ for both growth and development traits and morphology and physiology traits in figure 2. For the growth and development traits, the genetic variances of the first two out of six (1/3rd of the) eigenvectors of $\mathbf{G_{mf}}$ were fully accounted for by sexually concordant genetic variance (i.e. the first two eigenvectors of $\mathbf{G_C}$), and the third eigenvectors of $\mathbf{G_{mf}}$ was only partly explained by sexually concordant genetic variance (figure 2A). The remaining unexplained genetic variances in $\mathbf{G_{mf}}$'s third eigenvector is apparently sexually antagonistic, as indicated by the overabundance of genetic variance in the first eigenvector of $\mathbf{G_A}$ relative to the fourth eigenvector of $\mathbf{G_{mf}}$, and so on. By contrast, for the morphology and physiology traits only the first one out of eight (1/8th of the) eigenvectors of $\mathbf{G_{mf}}$ were fully accounted for by sexually concordant genetic variance (i.e. the first eigenvectors of $\mathbf{G_C}$), and all remaining eigenvectors of $\mathbf{G_{mf}}$ had some fraction of their genetic variances comprised of SA genetic variance ($\mathbf{G_A}$; figure 2B).

DISCUSSION

Mosses engage in scent-based fertilization in which female plants use specific VOCs to attract sperm-dispersing microarthropods, thereby increasing sexual reproduction [30–32]. Male mosses, in contrast, appear to produce fewer compounds, and in lower abundances, suggesting that VOC production may undergo sexually dimorphic selection [30,31]. Here, we used a multivariate approach based on field-collected, natural crosses to estimate the genetic architecture of variation in VOC production and life history traits in the moss *C. purpureus*. The study population contained genetic variance for all traits, consistent with previous studies of life history traits in other populations [34,53]. We found clear evidence for sexual dimorphism in the total number of masses produced and leaf length. Most traits have cross-sex correlations that were indistinguishable from zero, which would suggest that selection on one sex would elicit at most a modest response in the other sex. However, both Hansen's *d* and the simplified

eigentensor analysis showed that the multi-trait genetic (co)variance matrices, $\mathbf{G_f}$ and $\mathbf{G_m}$, were aligned, which would intuitively suggest that the multivariate pleiotropic constraints to the response to selection would be shared between the sexes. Still, the cross-trait cross-sex genetic (co)variance matrix (**B**) had asymmetric elements, indicating some opportunity for sex-limited responses to selection in spite of the putative multivariate genetic constraints indicated by the similarity between $\mathbf{G_f}$ and $\mathbf{G_m}$.

The constraint on the continued evolution of sexual dimorphism is typically evaluated by estimating the cross-sex correlations (r_{mf}) between homologous traits, and indeed the overall mixed r_{mf} values we found here are consistent with estimates from other populations of *C. purpureus* [50]. We found no relationship between r_{mf} and sexual dimorphism further supporting the inadequacy of r_{mf} as a metric of constraint. For example, total masses was sexually dimorphic but had a nearly zero r_{mf} while leaf length was similarly dimorphic and had a high non-zero r_{mf} (table 1). Additionally, juvenile growth form was not sexually dimorphic yet had a high non-zero r_{mf}. In other populations of *C. purpureus*, McDaniel [34] found a different relationship between dimorphism and r_{mf}, suggesting that this relationship may be highly population dependent. While diploid organisms may resolve constraints to sexual dimorphism via sex-specific dominance effects [14,54,55], conflict resolution in this haploid moss may be limited to alternative mechanisms such as sex-linkage or sex-chromosome mediated gene regulation. We suspect that a key factor explaining the mix of r_{mf} values in *C. purpureus* is the fact that females and males each have a large sex-limited chromosome (U: 3,450 genes and V: 3,411 genes, respectively) [56], where the U is passed from mother to daughter and the V from father to son, which could enable rapid resolution to sexual conflict. If so, this could mean that U- or V-linked variants may represent evolutionary changes aimed at resolving autosomal sexual conflict.

It is widely appreciated that single trait analyses, like r_{mf} may fail to capture the true underlying constraint on the evolution of sexual dimorphism. Indeed, estimates showing that male and female genetic (co)variance matrices are similar suggest that the response to

394 selection of one sex could be quite similar in the other in spite of the low cross-sex correlations
395 for individual homologous traits. Similar to findings in other studies [50,52,57–59], we found that
396 the overall genetic (co)variance structure was similar between males and females (table 2).
397 Despite similar sex specific covariance matrices, there are some observable differences,
398 including the negative covariance of leaf length and total masses in males but not females, and
399 reproductive effort and leaf length positively covary in females but not in males (figure 1). Many
400 of the most differentiated covariances involved leaf traits and relative reproductive effort with
401 VOC production in mature plants. The fact that many traits show cross-trait covariances that are
402 sexually dimorphic suggests that genetic control is both highly pleiotropic (between traits) and
403 potentially involves strong epistatic interactions with loci on the U and V sex chromosomes. In
404 addition, this suggests that similar patterns of selection acting on males or females could
405 generate different phenotypic responses, potentially increasing or decreasing the population-
406 level sexual dimorphism.
407 Intuitively, it would make sense that similarity between $\mathbf{G_f}$ and $\mathbf{G_m}$ would impose genetic
408 constraint. However, Cheng and Houle [20] demonstrated that similarity in male and female
409 covariance matrices coupled with some degree of **B** matrix asymmetry suggests a greater
410 opportunity for sexual dimorphism in response to sexually concordant selection than to sexually
411 antagonistic selection. Thus, our estimates of the proportion of standing genetic variation that
412 could respond to sexually antagonistic selection represent lower bounds for the potential
413 sexually dimorphic response, as further sexual dimorphism could evolve in response to sexually
414 concordant selection. We therefore base our findings regarding multivariate genetic constraint
415 on the estimated proportions of asymmetry and symmetry on our **B** matrix analysis [51,52].
416 Though **B** was largely symmetrical, indicating multivariate constraints to sexual
417 dimorphism, a portion of the **B** matrix was asymmetric in both trait categories (growth and
418 development and morphology and physiology). If the **B** matrix were completely symmetrical, the
419 response to selection on males would be manifest in both the male and female offspring of the

following generation. By contrast, asymmetry in the off diagonals of the **B** matrix means that the multivariate responses to selection between males and females can be different [29,50,59,60]. The asymmetry in **B** likely results from sex-biased gene regulation mediated by epistatic interactions between autosomal variants and the U and V sex chromosomes (possibly also mediated by epigenetic factors; see Wang et al. [61]). There seems to be at least a putative difference between the growth and development traits and the morphology physiology traits in the degree of **B** asymmetry (table 2), which is also visually apparent in figure 1. The levels of **B** asymmetry that we find in the growth and development traits and morphology physiology traits is toward the lower and upper end, respectively, of the range of estimates among populations of *Drosophila serrata* [51], which ranged from ~15-30% (table 2). This possibly suggests a richer history of sex-specific and/or sexually antagonistic selection in morphology and physiology traits relative to growth and development traits, triggering the evolution of resolved genetic constraints.

An analysis of the degree of multivariate sexually antagonistic genetic variation in $\mathbf{G}_{mf}$ provides insight to the capacity for further response to sexually antagonistic selection [20,52]. The overall percentages of sexually antagonistic genetic variance were estimated with wide, highly overlapping CIs between our two trait categories (table 2). However, the eigenvector-specific analysis showed a greater proportion of sexually antagonistic genetic variance comprising the eigenvectors of $\mathbf{G}_{mf}$ in morphology and physiology traits relative to the growth and development traits. Further, that sexually antagonistic genetic variance was dispersed across proportionally more of the eigenvectors relative to that exhibited by the growth and development traits (figure 2). Indeed, 25-35% of the multivariate genetic variance in our population was sexually antagonistic (table 2), considerably more than, for example, the multivariate genetic architecture of wing morphology in *D. melanogaster* (4.32% sexually antagonistic genetic variance [52]). Thus, our morphology and physiology traits may possess a greater opportunity to respond to sexually antagonistic selection than the growth and

development traits, echoing the greater proportion of the **B** matrix that was found to be asymmetric relative to that of growth and development traits (figure 1, table 2).

The rich bouquet of VOCs produced by this population may contribute to variation in attracting sperm-dispersing arthropods, with potentially major fitness consequences. Both females and males contain genetic variation for VOC production, but the structure of covariation in the sexes is sufficiently different such that sex-specific coevolution between the moss scents and arthropod behaviors could play a major role in the maintenance of genetic variation for fitness in natural populations of *C. purpureus*. The complexity of the underlying genetic architecture also highlights the potential for scent-based fertilization to contribute to pre-zygotic speciation barriers in mosses, much like the role pollination plays in angiosperms. For example, mosses may evolve suites of VOCs which match the preferences of the local mesofauna. Odor-mediated fertilization could promote the evolution of pre-zygotic isolation if moss VOCs elicit species-specific responses from sperm-dispersing microarthropods or other members of these communities. It is possible that the interaction involves additional microbial partners upon which the mesofauna feed – indeed, mosses appear to host diverse sex- and species-specific microbiomes [62–64]. Collectively these results highlight how ecological interactions may shape the evolution of sexual dimorphism [65,66], which may in turn contribute to the maintenance of genetic variation in fitness and the evolution of reproductive isolation.

Data accessibility

Data and scripts to reproduce the results of the study are available on the Dryad Repository: https://doi.org/10.5061/dryad.59zw3r266.

Contributions

L.M.K. led the study. L.M.K, T.N.R, and S.F.M designed the study. L.M.K. collected the data, with assistance from T.N.C., T.K., A.J.J., D.N.S., and C.T.C. for the life history traits, and E.T.G. and S.K. for the PTR-TOF-MS data. L.M.K. and K.G. performed the statistical analyses where

K.G. supervised statistical analyses and interpretation. L.M.K. and S.F.M. wrote the manuscript, E.T.G. and K.G. contributed to writing the methods and K.G. contributed to writing results and editing the manuscript.

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Figure Captions

Figure 1. Genetic correlations (Gmf) among traits within and between males and females represented by ellipses. A narrow ellipse is representative of a stronger correlation while a wider ellipse depicts a weaker correlation. A represents the genetic correlations between growth and developmental traits whereas B represents the correlations between morphology and physiology.

Figure 2. A comparison of the genetic variance of Gmf against the concordant and antagonistic subspaces. The height of each bar represents the estimated genetic variance for each eigenvector while the error bars show the 95% HPD. Plot A (6 dimensions) represents the growth and development traits, and plot B (8 dimensions) represents the morphology and physiology traits.

Table 1. Estimates of sex specific genetic variance and associated 95% HPD intervals and cross- sex correlations (rmf) and associated 95% HPD intervals. The degree sexual dimorphism was calculated as the difference between point estimates of male and female posterior means (male – female). A negative value for sexual dimorphism suggests the females have a larger posterior mean. All traits with an "*" are sexually dimorphic ($p < 0.05$).

Table 2. Summary table with estimates and corresponding 95% HPD intervals and p- values where applicable. Estimates include comparisons between Gm and Gf (Hansen's difference d and simplified eigentensor analysis), asymmetry and symmetry of B, and proportion of antagonistic and concordant subspace relative to the total genetic variance in Gmf.

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