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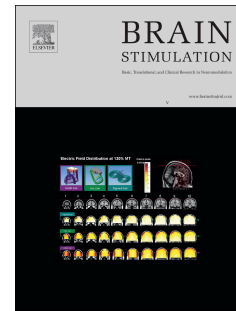
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1 **False positives associated with responder/non-responder analyses**
2 **based on motor evoked potentials**

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8

9 **Keywords:** variability, MEP, TMS, plasticity, corticospinal excitability, responders

10

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21 Abstract

22 Background: A trend in the non-invasive brain stimulation literature is to assess the outcome of an
23 intervention using a responder analysis whereby participants are di- or trichotomised in order that they
24 may be classified as either responders or non-responders.

25 Objective: Examine the extent of the Type I error in motor evoked potential (MEP) data subjected to
26 responder analyses.

27 Methods: Seven sets of 30 MEPs were recorded from the first dorsal interosseous muscle in 52 healthy
28 volunteers. Four classification techniques were used classify the participants as responders or non-
29 responders: (1) the two-step cluster analysis, (2) Dichotomised thresholding, (3) relative method and
30 (4) baseline variance method.

31 Results: Despite the lack of any intervention, a significant number of participants were classified as
32 responders (21-71%).

33 Conclusion: This study highlights the very large Type I error associated with dichotomising
34 continuous variables such as the TMS MEP.

35 Introduction

36 Similar to many other interventions, the efficacy of non-invasive brain stimulation (NIBS) is limited to
37 a subset of the population and it is important to better understand what proportion of participants
38 might respond. A recent trend in the NIBS literature is to use a responder analysis to classify
39 participants as responders or non-responders following an intervention. This simplifies the statistical
40 analysis, interpretation and presentation of results [1]. In the NIBS literature, this classification is
41 typically performed by di- or trichotomising the motor evoked potential (MEP) produced in response
42 to transcranial magnetic stimulation (TMS) as this is considered a surrogate marker of neuroplasticity
43 [2].

44 Pellegrini, et al. 2018 [3] recently conducted a systematic review of responder analyses in NIBS and
45 concluded that they can effectively identify subgroups based on response patterns, and used to
46 estimate the proportion of participants who might respond to the intervention. However, they also
47 noted a lack of consistency and consensus in the methods by which the data are quantified.
48 Furthermore, they highlighted that many studies in the NIBS literature lack a control group. As a
49 result, the effect of natural variability of the MEP is not accounted for with these analyses. The MEP
50 magnitude has considerable trial-to-trial variability and drift over time, which arise due to controllable
51 and uncontrollable factors of physiological (e.g. cortical rhythms, arousal, etc.) and non-physiological
52 (e.g. TMS coil placement and/or movement) origin [4, 5].

53 Responder analyses methods gained popularity in the early 2000s in the clinical medicine and
54 psychology literature primarily as a means to establish proportions of responders in drug trials and in
55 marketing studies [6-8]. However, these methods were then criticised by methodologists who
56 questioned the validity of dichotomising (or trichotomising) continuous variables. They noted in
57 particular that inferences made from such analyses are susceptible to large Type I error (false
58 positives) that can lead to erroneous conclusions [1, 6, 9-19]. The aim of the present study was to
59 examine the extent of the Type I error in MEP data that are subjected to different types responder
60 analyses.

61

62 **Methods**

63 *Experimental procedures*

64 Fifty-two healthy participants, without contraindication to TMS and no history of neurological
65 psychiatric disorder, participated in the study (20 ± 2 y, range 18-25, 35 female). Participants visited
66 the laboratory once for ~1 h, during which MEPs were recorded from the first dorsal interosseus
67 (FDI). Participants sat comfortably and were instructed to relax both the hand and arm, and to keep
68 their eyes open for the duration of the experiment. To facilitate this instruction throughout the
69 experiment, interactive feedback of FDI muscle activity was provided on a computer monitor. TMS
70 was delivered through a 90 mm figure-of-8 coil (type: batwing; type no. 15411) using a Magstim
71 Rapid² stimulator (Magstim Ltd, Dyfed, United Kingdom). Coil position and orientation were
72 monitored with frameless stereotaxy (BrainSight 2, Rogue Research Inc, Montreal, Canada). The
73 stimulation intensity required to evoke 1 mV (SI_{1mV}) peak-to-peak MEPs (MEP_{pp}) was determined by
74 adjusting the intensity until the mean of 30 stimuli produced a 1 mV MEP_{pp} (calibration data set in
75 Figure 1A). Next, seven sets of 30 MEPs were recorded with a 4 s inter-stimulus interval and 2 min
76 rest between sets. The first set was deemed a baseline to which the remaining 6 data sets would be
77 compared. Figure 1A summarises the experimental protocol.

78 *Statistical Analysis*

79 The MEP_{pp} amplitude was extracted between 20-50 ms after stimulation and averaged across all
80 stimuli within a set. The mean MEP_{pp} for each set was then used for statistical analysis and
81 classification either: (1) without any further processing; or (2) after normalisation to the mean MEP_{pp}
82 of the baseline set (B), the 'grand average (GA) method'. Therefore, each classification method was
83 performed twice on the same data, either the absolute mean MEP_{pp} amplitudes for each set, or the
84 normalised GA data.

85 Before classification, the continuous data was analysed using a repeated measures analysis of variance
86 (RM-ANOVA) across sets for the mean absolute MEP_{pp} values. Subsequently, the participants were
87 classified using the four common methods found in the NIBS literature. Following classification, a
88 mixed RM-ANOVA was performed on the absolute MEP_{pp} data with the within-factor 'set' and

89 between-subjects factor 'group' (i.e. the result of the classification method). In addition, a one-way
90 RM-ANOVA was performed for each group individually on the absolute MEP_{pp} data to classify
91 groups of participants as either:

- 92 • (+) responders: significant increase in MEP_{pp} across set
- 93 • (-) responders: significant decrease in MEP_{pp} across set
- 94 • (0) responders or non-responders: no significant change in MEP_{pp} across set

95 If Mauchly's Test of Sphericity indicated that the assumption of sphericity had been violated, a
96 Greenhouse-Geisser correction (GG) was performed. All statistical tests were performed using SPSS,
97 with significance accepted at $p < 0.05$.

98 *Responder Analysis Methods*

99 1) *Two-step cluster analysis*: This SPSS method uses a two-step clustering approach that allows
100 automatic detection of the optimal number of clusters. In the first step all cases are scanned and
101 pre-clustered based on a predefined distance criterion (e.g. squared Euclidian distance or log-
102 likelihood) that specifies either the difference or similarity between cases. In the second step,
103 the algorithm uses agglomerative hierarchical clustering to merge the sub clusters resulting
104 from the first step into a smaller number of clusters. In the present study we allowed the
105 algorithm to automatically determine the number of clusters rather than specifying two or
106 three clusters. This is a commonly used method in NIBS literature [20-26].

107 2) *Dichotomised thresholding*: This method separates data into two groups based on a predefined
108 threshold. For GA data, participants were categorised using the mean GA of sets (in our case
109 sets T1-T6). Participants were then classified as negative responders for mean GA < 1 and
110 positive responders for mean GA > 1. This analysis was also performed on absolute MEP_{pp}
111 data. With absolute MEP_{pp} data this method can be applied either on a group level or
112 individually. For the group level analysis, the mean MEP_{pp} amplitude across all participants
113 was chosen as the threshold (1.35 mV in this study). For the individual analysis, the threshold
114 is set to the mean MEP_{pp} of the baseline set for each participant individually. Next, each
115 participant is classified as a positive responder if the mean MEP_{pp} across T1-T6 is greater than

116 the threshold and a negative responder if the mean MEP_{pp} across T1-T6 is less than the
117 threshold. Dichotomised thresholding is a common method of subgrouping normalised MEP
118 data [22, 24-33].

119 3) *Relative method*: This method is used to classify participants into three groups based on a
120 predefined percent change from baseline threshold. This method has been used in several
121 studies to trichotomise participants using a threshold of 10% [23, 34], 15% [35], 20% [20] or
122 50% [36]. In the present study we used a conservative approach by choosing 20% change
123 from baseline as the threshold. For the GA data, participants are classified as negative
124 responders for mean GA across sets T1-T6 < 0.8 , positive responders for mean GA > 1.2 and
125 non-responders between 0.8-1.2. Likewise for the absolute MEP_{pp} data the threshold was 1.35
126 ± 0.27 mV as for the collected data the group mean of the baseline set B was 1.35 mV. This
127 procedure was also performed on an individual level, in which case the threshold was
128 individually determined based on the mean MEP_{pp} amplitude of set B.

129 4) *Baseline variance method*: In this method participants are trichotomised based on the variance
130 of the baseline measure. For the GA data, the standard error (SE) of the GA of the baseline set
131 was 0.14 across all participants. Therefore, a participant was classified as a (-) or (+)
132 responder if the mean GA across sets T1-T6 was smaller or greater than 1.27 (95% confidence
133 limit (CL) 1 ± 0.27) and a non-responder otherwise. Similarly, for MEP_{pp} data the SE of the
134 baseline set was 0.17 across all participants (95% CL 1.35 ± 0.36 mV) and therefore a
135 participant was a (+) responder when above this upper limit, a (-) when below the lower limit
136 or a non-responder otherwise. The same analysis was also performed on the level of each
137 individual, i.e. the CL of the baseline set was determined individually to assign the participant
138 to the correct group. This method has been used in several studies [28, 33, 37-41].

139

140 **Results**

141 A one-way RM-ANOVA applied across all seven data sets (B-T6) before dichotomisation revealed
142 neither a significant difference in mean MEP_{pp} amplitude across these data sets ($F_{(4,76,242.75)} = 1.27^{GG}$,
143 $p=0.28$) nor in GA ($F_{(4,74,241.73)} = 1.31^{GG}$ $p=0.26$; Figure 1B).

144 The results for the subgrouping methods are presented in Table 1 and for the group level analysis
145 visualized in Figure 1C. The SPSS two-step cluster analysis determined two clusters to best separate
146 the data. For the MEP_{pp} data 11 participants (~21%) were classified as responders, showing a
147 significant increase in MEP_{pp} ($p<0.01$) across time, and 41 participants (~79%) were classified as non-
148 responders ($p=0.96$). The same groups were identified using the GA data but with 19 responders
149 ($p<0.01$) and 33 non-responders ($p=0.22$). The MEP_{pp} and GA across time for each group is illustrated
150 in Figure 1C.

151 Using the dichotomised thresholding method on MEP_{pp} data and a group level, 33 participants (63%)
152 were classified as (+) responders ($p<0.01$) and 19 participants (37%) as non-responders ($p=0.88$). For
153 the GA data, 28 participants (54%) were classified as (+) responders ($GA > 1$, $p<0.01$) and 24
154 participants (46%) were classified as (-) responders ($GA < 1$, $p=0.01$) (Figure 1D).

155 The relative and baseline variance methods produced similar proportions of responders when
156 performed irrespective of the group or individual level analysis. Generally, more participants were
157 classified as non-responders for the GA data (40-58%) than the MEP_{pp} data (29-52%). Moreover, the
158 baseline variance method resulted in more non-responders (46-58%) than the relative method (29-
159 40%).

160

161 **Discussion**

162 The present study followed a typical intervention design where TMS MEP data are collected at
163 baseline and then again at pre-defined times following the intervention. However, in the present study
164 the participants were not exposed to an intervention. Therefore, subject to normal MEP variability, the
165 'post-intervention' data sets would not be expected to be different from baseline. As expected,
166 parametric statistics performed on this continuous data set revealed no significant difference with time.
167 However, when the data were subjected to responder analysis between 21-71% of the participants were
168 classified as responders, thus revealing a large number of false positives.

169 The responder analysis has been used throughout clinical medicine and psychology literature because
170 it simplifies the analysis and interpretation of experimental results; with proponents of the analysis
171 highlighting its usefulness in clinical decision making [7]. However, methodologists have argued for
172 more than two decades that the dichotomisation of continuous variables is not valid for hypothesis
173 testing [1, 9-14, 16-18]. The dichotomisation of continuous variables results in significant loss of
174 information (~35-50% depending on the distribution of the data), reduced power of the statistical tests,
175 high probability of Type I error, biased parameter estimates and erroneously small variances (for
176 detailed discussion see: [1, 13, 16]).

177 The specific objective of the present study was to investigate the Type I error associated with
178 responder analyses when MEP data are used to classify participants. In general, we observed
179 substantial Type I errors with all of the responder analyses methods. Our results suggest that at best,
180 20% of the participants who have been classified as responders will have been classified erroneously.
181 It may be valid to use a responder analysis to compare an intervention with a control group, but the
182 specific response rates may be over-estimated.

183

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186 Clark for valuable discussions with respect to the responder analysis and statistical processing.

187 Conflict of Interest:

188 We have no conflicts of interest to declare.

189 Ethical approval:

190 The study was carried out in accordance with The Code of Ethics of the World Medical Association
191 (Declaration of Helsinki) and informed consent was obtained from all participants recruited to the
192 study. Ethical approval for the study was granted from the University of Birmingham's Science,
193 Technology, Engineering and Mathematics ethics committee (ERN_13-0701).

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196 not-for-profit sectors.

197

198 **References**

- 199 [1] Altman DG, Royston P. The cost of dichotomising continuous variables. *Bmj*
200 2006;332(7549):1080.
- 201 [2] Rossini PM, Burke D, Chen R, Cohen LG, Daskalakis Z, Di Iorio R, et al. Non-invasive
202 electrical and magnetic stimulation of the brain, spinal cord, roots and peripheral nerves: Basic
203 principles and procedures for routine clinical and research application. An updated report from an
204 I.F.C.N. Committee. *Clin Neurophysiol* 2015;126(6):1071-107.
- 205 [3] Pellegrini M, Zoghi M, Jaberzadeh S. Cluster analysis and subgrouping to investigate inter-
206 individual variability to non-invasive brain stimulation: a systematic review. *Reviews in the*
207 *neurosciences* 2018.
- 208 [4] Schmidt S, Bathe-Peters R, Fleischmann R, Ronnefarth M, Scholz M, Brandt SA.
209 Nonphysiological factors in navigated TMS studies; confounding covariates and valid intracortical
210 estimates. *Hum Brain Mapp* 2015;36(1):40-9.
- 211 [5] Kiers L, Cros D, Chiappa KH, Fang J. Variability of Motor Potentials-Evoked by Transcranial
212 Magnetic Stimulation. *Electroencephalography and Clinical Neurophysiology* 1993;89(6):415-23.
- 213 [6] Senn S, Julious S. Measurement in clinical trials: A neglected issue for statisticians? *Statistics*
214 *in Medicine* 2009;28(26):3189-209.
- 215 [7] Snapinn SM, Jiang Q. Responder analyses and the assessment of a clinically relevant
216 treatment effect. *Trials* 2007;8(1):31.
- 217 [8] Iacobucci D, Popovich DL, Bakamitsos GA, Posavac SS, Kardes FR. Three Essential
218 Analytical Techniques for the Behavioral Marketing Researcher: Median Splits, Mean-Centering, and
219 Mediation Analysis. *Foundations and Trends® in Marketing* 2015;9(2):83-174.
- 220 [9] Weinberg CR. How bad is categorization? *Epidemiology* 1995;6(4):345-7.
- 221 [10] Senn S. Disappointing dichotomies. *Pharm Stat* 2003;2(4):239-40.
- 222 [11] Royston P, Altman DG, Sauerbrei W. Dichotomizing continuous predictors in multiple
223 regression: a bad idea. *Stat Med* 2006;25(1):127-41.
- 224 [12] Metze K. Dichotomization of continuous data--a pitfall in prognostic factor studies. *Pathol Res*
225 *Pract* 2008;204(3):213-4.
- 226 [13] Maxwell SE, Delaney HD. Bivariate Median Splits and Spurious Statistical Significance.
227 *Psychol Bull* 1993;113(1):181-90.
- 228 [14] MacCallum RC, Zhang S, Preacher KJ, Rucker DD. On the practice of dichotomization of
229 quantitative variables. *Psychol Methods* 2002;7(1):19-40.
- 230 [15] Lewis JA. In defence of the dichotomy. *Pharm Stat* 2004;3(2):77-9.
- 231 [16] Fedorov V, Mannino F, Zhang R. Consequences of dichotomization. *Pharm Stat* 2009;8(1):50-
232 61.
- 233 [17] DeCoster J, Iselin AM, Gallucci M. A conceptual and empirical examination of justifications
234 for dichotomization. *Psychol Methods* 2009;14(4):349-66.
- 235 [18] Cohen J. The cost of dichotomization. *Applied Psychological Measurement* 1983;7(3):249-53.
- 236 [19] Julie R. Irwin, McClelland GH. Negative Consequences of Dichotomizing Continuous
237 Predictor Variables. *Journal of Marketing Research* 2003;40(3):366-71.
- 238 [20] Chew T, Ho KA, Loo CK. Inter- and Intra-individual Variability in Response to Transcranial
239 Direct Current Stimulation (tDCS) at Varying Current Intensities. *Brain stimulation* 2015.
- 240 [21] López-Alonso V, Cheeran B, Fernández-del-Olmo M. Relationship between non-invasive
241 brain stimulation-induced plasticity and capacity for motor learning. *Brain stimulation*
242 2015;8(6):1209-19.
- 243 [22] Lopez-Alonso V, Cheeran B, Rio-Rodriguez D, Fernandez-Del-Olmo M. Inter-individual
244 variability in response to non-invasive brain stimulation paradigms. *Brain stimulation* 2014;7(3):372-
245 80.
- 246 [23] Puri R, Hinder MR, Canty AJ, Summers JJ. Facilitatory non-invasive brain stimulation in
247 older adults: the effect of stimulation type and duration on the induction of motor cortex plasticity.
248 *Experimental brain research* 2016;234(12):3411-23.
- 249 [24] Puri R, Hinder MR, Fujiyama H, Gomez R, Carson RG, Summers JJ. Duration-dependent
250 effects of the BDNF Val66Met polymorphism on anodal tDCS induced motor cortex plasticity in older
251 adults: a group and individual perspective. *Frontiers in aging neuroscience* 2015;7:107.

- 252 [25] Strube W, Bunse T, Nitsche MA, Nikolaeva A, Palm U, Padberg F, et al. Bidirectional
253 variability in motor cortex excitability modulation following 1 mA transcranial direct current
254 stimulation in healthy participants. *Physiol Rep* 2016;4(15).
- 255 [26] Wiethoff S, Hamada M, Rothwell JC. Variability in response to transcranial direct current
256 stimulation of the motor cortex. *Brain stimulation* 2014;7(3):468-75.
- 257 [27] Goldsworthy MR, Vallence AM, Yang R, Pitcher JB, Ridding MC. Combined transcranial
258 alternating current stimulation and continuous theta burst stimulation: a novel approach for
259 neuroplasticity induction. *Eur J Neurosci* 2016;43(4):572-9.
- 260 [28] Hamada M, Murase N, Hasan A, Balaratnam M, Rothwell JC. The role of interneuron
261 networks in driving human motor cortical plasticity. *Cerebral cortex* 2013;23(7):1593-605.
- 262 [29] Hinder MR, Goss EL, Fujiyama H, Canty AJ, Garry MI, Rodger J, et al. Inter- and Intra-
263 individual variability following intermittent theta burst stimulation: implications for rehabilitation and
264 recovery. *Brain stimulation* 2014;7(3):365-71.
- 265 [30] Labruna L, Jamil A, Fresnoza S, Batsikadze G, Kuo MF, Vanderschelden B, et al. Efficacy of
266 Anodal Transcranial Direct Current Stimulation is Related to Sensitivity to Transcranial Magnetic
267 Stimulation. *Brain stimulation* 2016;9(1):8-15.
- 268 [31] Lopez-Alonso V, Fernandez-Del-Olmo M, Costantini A, Gonzalez-Henriquez JJ, Cheeran B.
269 Intra-individual variability in the response to anodal transcranial direct current stimulation. *Clin*
270 *Neurophysiol* 2015.
- 271 [32] Muller-Dahlhaus JF, Orekhov Y, Liu Y, Ziemann U. Interindividual variability and age-
272 dependency of motor cortical plasticity induced by paired associative stimulation. *Experimental brain*
273 *research* 2008;187(3):467-75.
- 274 [33] Nakamura K, Groiss SJ, Hamada M, Enomoto H, Kadowaki S, Abe M, et al. Variability in
275 Response to Quadripulse Stimulation of the Motor Cortex. *Brain stimulation* 2016;9(6):859-66.
- 276 [34] Muller-Dahlhaus F, Lucke C, Lu MK, Arai N, Fuhl A, Herrmann E, et al. Augmenting LTP-
277 Like Plasticity in Human Motor Cortex by Spaced Paired Associative Stimulation. *Plos One*
278 2015;10(6):e0131020.
- 279 [35] Nettekoven C, Volz LJ, Leimbach M, Pool EM, Rehme AK, Eickhoff SB, et al. Inter-
280 individual variability in cortical excitability and motor network connectivity following multiple blocks
281 of rTMS. *NeuroImage* 2015;118:209-18.
- 282 [36] Strube W, Bunse T, Malchow B, Hasan A. Efficacy and interindividual variability in motor-
283 cortex plasticity following anodal tDCS and paired-associative stimulation. *Neural plasticity*
284 2015;2015:530423.
- 285 [37] Ammann C, Lindquist MA, Celnik PA. Response variability of different anodal transcranial
286 direct current stimulation intensities across multiple sessions. *Brain stimulation* 2017;10(4):757-63.
- 287 [38] Hanajima R, Tanaka N, Tsutsumi R, Enomoto H, Abe M, Nakamura K, et al. The effect of age
288 on the homotopic motor cortical long-term potentiation-like effect induced by quadripulse stimulation.
289 *Experimental brain research* 2017;235(7):2103-8.
- 290 [39] Simeoni S, Hannah R, Sato D, Kawakami M, Rothwell J, Simeoni S, et al. Effects of
291 Quadripulse Stimulation on Human Motor Cortex Excitability: A Replication Study. *Brain stimulation*
292 2016;9(1):148-50.
- 293 [40] Tremblay S, Hannah R, Rawji V, Rothwell JC. Modulation of iTBS after-effects via
294 concurrent directional TDCS: A proof of principle study. *Brain stimulation* 2017;10(4):744-7.
- 295 [41] Tremblay S, Larochelle-Brunet F, Lafleur LP, El Mouderrib S, Lepage JF, Theoret H.
296 Systematic assessment of duration and intensity of anodal transcranial direct current stimulation on
297 primary motor cortex excitability. *Eur J Neurosci* 2016;44(5):2184-90.

299 **Figure/Table Legends**

300 **Figure 1:** Responder/non-responder analysis across TMS MEP testing sets. (A) Seven sets of 30
301 MEPs were acquired at a stimulation intensity selected to producing a mean 1 mV peak-to-peak MEP
302 amplitude (mean SI_{1mV} : $56 \pm 10\%$ of maximum stimulator output). The first set was considered the
303 baseline to which the remaining six sets would be compared. (B) MEP_{pp} amplitude across all
304 participants and all sets. No effect of set on MEP_{pp} amplitude observed for these data. (C) MEP_{pp}
305 amplitude is shown across each of the seven data sets, with the participants di- or tricotomised using a
306 two-step cluster analysis, dichotomised thresholding, relative threshold method or baseline variance
307 method on a group level. In this way participants are classified as either (+) responders (light grey
308 lines), showing an increase in MEP_{pp} amplitude compared to baseline, (0)- or non-responders (grey
309 lines), no change in MEP_{pp} amplitude across SET, or (-) responders (black lines), a decrease in
310 absolute MEP_{pp} across SET. The left column presents results when the classification was based on
311 absolute MEP_{pp} data, the right column when based on GA data. All data are presented as Mean \pm S.D.
312 The number of participants for each group can be found in Table 1.

313 **Table 1:** Overview of results for subgrouping participants according to four methods for both
314 normalised grand average (GA) data as well as non-normalised 'raw' MEP_{pp} data: (1). SPSS Two-Step
315 Cluster analysis; (2) Relative % change with respect to baseline; (3) Dichotomised thresholding: a
316 predefined fixed threshold; and (4) Change relative to the variance of the baseline set. A subgroup of
317 participants is classified as positive responders (+) or negative responders (-), when there is a
318 significant increase or decrease across SET respectively. Non-responders (0) are those participants in
319 the group with no significant change in MEP_{pp} amplitude across SET. For some methods participants
320 were subgrouped both on a threshold defined on an individual (Indv) basis as well as on a group (Gr)
321 level. The %0 column highlights the proportion of non-responders. Results are shown with analysis
322 performed on normalised grand average (GA) data and non-normalised absolute MEP_{pp} data.

323

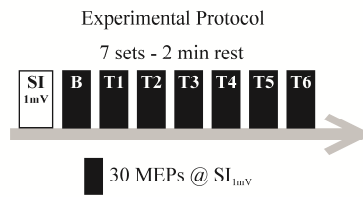
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325

Normalised GA data														
Subgrouping Method	# Participants				Mixed RM-ANOVA			OneWay RM-ANOVA						
	+	0	-	%0				+	0		-			
Two Step Cluster		19	33	-	63%	SET: SET×GROUP:	$F_{(4.83,241.71)} = 3.43^{GG}$ $F_{(4.83,241.71)} = 8.40^{GG}$	$p < 0.01$ $p < 0.01$	$F_{(3.66,65.93)} = 5.97^{GG}$	$p < 0.01$	$F_{(5.02,160.76)} = 1.65^{GG}$	$p = 0.15$	-	-
Threshold Dichotomisation		28	-	24	-	SET: SET×GROUP:	$F_{(4.88,243.73)} = 1.05^{GG}$ $F_{(4.88,243.73)} = 8.14^{GG}$	$p = 0.39$ $p < 0.01$	$F_{(3.96,106.90)} = 6.33^{GG}$	$p < 0.01$	-	-	$F_{(6,138)} = 2.78$	$p = 0.01$
Relative		20	21	11	40%	SET: SET×GROUP:	$F_{(4.66,228.43)} = 0.49^{GG}$ $F_{(9.32,228.43)} = 5.63^{GG}$	$p = 0.77$ $p < 0.01$	$F_{(3.69,70.22)} = 5.91^{GG}$	$p < 0.01$	$F_{(4.41,88.25)} = 0.64^{GG}$	$p = 0.65$	$F_{(6,60)} = 4.59$	$p < 0.01$
Baseline Variance	Gr	15	27	10	52%	SET: SET×GROUP:	$F_{(4.63,226.73)} = 0.97^{GG}$ $F_{(9.25,226.73)} = 6.08^{GG}$	$p = 0.43$ $p < 0.01$	$F_{(6,84)} = 6.59$	$p < 0.01$	$F_{(4.59,119.21)} = 0.52^{GG}$	$p = 0.74$	$F_{(6,54)} = 4.29$	$p < 0.01$
	Indv	13	30	9	58%	SET: SET×GROUP:	$F_{(4.57,223.80)} = 1.24^{GG}$ $F_{(9.14,223.80)} = 6.59^{GG}$	$p = 0.29$ $p < 0.01$	$F_{(3.11,37.37)} = 6.68^{GG}$	$p < 0.01$	$F_{(4.56,132.17)} = 0.48^{GG}$	$p = 0.77$	$F_{(6,48)} = 4.58$	$p = 0.01$
Non-normalised MEP _{DD} data														
Two Step Cluster		11	41	-	79%	SET: SET×GROUP:	$F_{(6,300)} = 4.74$ $F_{(6,300)} = 4.96$	$p < 0.01$ $p < 0.01$	$F_{(6,60)} = 4.50$	$p < 0.01$	$F_{(6,240)} = 0.26$	$p = 0.96$	-	-
Threshold Dichotomisation	Gr	33	19	-	37%	SET: SET×GROUP:	$F_{(6,300)} = 3.23$ $F_{(6,300)} = 6.69$	$p < 0.01$ $p < 0.01$	$F_{(3.65,65.65)} = 5.80^{GG}$	$p < 0.01$	$F_{(6,192)} = 0.88$	$p = 0.51$	-	-
	Indv	24	-	28	-	SET: SET×GROUP:	$F_{(4.87,243.27)} = 1.06^{GG}$ $F_{(4.87,243.27)} = 7.44^{GG}$	$p = 0.38$ $p < 0.01$	$F_{(3.81,102.80)} = 5.80^{GG}$	$p < 0.01$	-	-	$F_{(6,138)} = 2.57$	$p = 0.02$
Relative	Gr	16	15	21	29%	SET: SET×GROUP:	$F_{(6,294)} = 2.12$ $F_{(12,294)} = 5.23$	$p = 0.05$ $p < 0.01$	$F_{(3.62,52.85)} = 5.00^{GG}$	$p < 0.01$	$F_{(6,84)} = 2.43$	$p = 0.03$	$F_{(6,120)} = 2.91$	$p = 0.01$
	Indv	17	19	16	37%	SET: SET×GROUP:	$F_{(4.73,231.96)} = 1.47^{GG}$ $F_{(9.47,231.96)} = 6.63^{GG}$	$p = 0.20$ $p < 0.01$	$F_{(3.41,54.60)} = 6.44^{GG}$	$p < 0.01$	$F_{(6,108)} = 1.70$	$p = 0.13$	$F_{(6,90)} = 4.13$	$p < 0.01$
Baseline Variance	Gr	12	27	13	52%	SET: SET×GROUP:	$F_{(4.75,232.84)} = 2.16^{GG}$ $F_{(9.50,232.84)} = 6.95^{GG}$	$p = 0.06$ $p < 0.01$	$F_{(3.10,34.09)} = 6.32^{GG}$	$p < 0.01$	$F_{(6,156)} = 1.51$	$p = 0.18$	$F_{(6,72)} = 4.77$	$p < 0.01$
	Indv	13	24	15	46%	SET: SET×GROUP:	$F_{(4.79,234.73)} = 2.49^{GG}$ $F_{(9.58,234.73)} = 6.08^{GG}$	$p = 0.03$ $p < 0.01$	$F_{(3.11,37.37)} = 6.68^{GG}$	$p < 0.01$	$F_{(6,138)} = 0.95$	$p = 0.36$	$F_{(6,84)} = 3.41$	$p < 0.01$

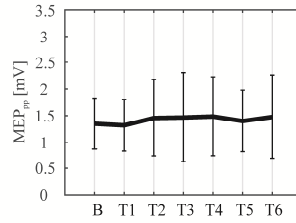
Table 1

A)



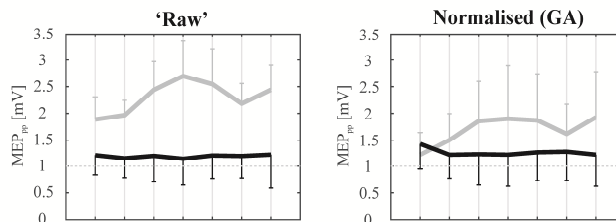
B)

Continuous

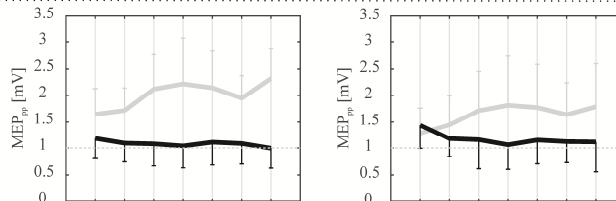


C)

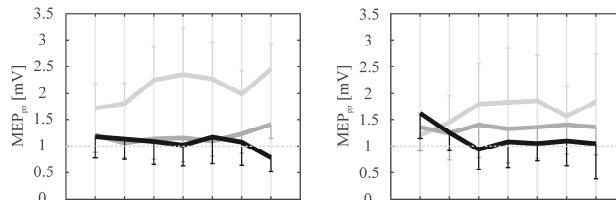
TwoStep Cluster



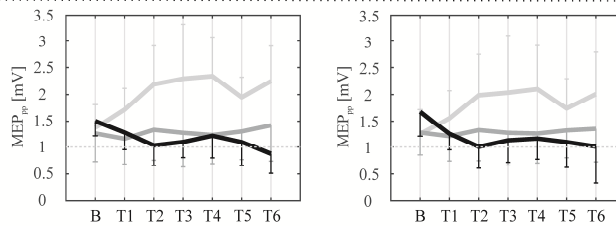
Dichotomised Thresholding



Relative Method



Baseline Variance Method



— (+) Responders
 — (0) Non-Responders
 — (-) Responders